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# Gender-matching School Effects on Girls' Cognitive and Non-cognitive Performance - Empirical Evidence from South Korea. <br> Cho, Seo-Young <br> MACIE, Philipps-Universität Marburg 

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Gender-matching School Effects on GirlsôCognitive and Non-cognitive Performance ð Empirical Evidence from South Korea

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#### Abstract

This paper investigates whether gender-matching school environments can improve girlsôinterest and motivation in science. Using the PISA data from South Korea, the empirical results of this paper show that gender-matching school environments have the most positive effect on the attitudinal development of girls who are at the highest quartile of science studies. By attending an all-girls school and being taught by a female science teacher, high-performing girls become as motivated and interested in pursuing science studies and careers as boys. Additionally, female teachers generally have a positive effect on developing girlsô competitive attitudes regardless of their cognitive performance. However, the effect of single-sex schooling is heterogeneous across different student groups. While maintaining its positive effect on high and low-performing girls, attending an all-girls school can be detrimental to the non-cognitive development of median girls. These findings corroborate that gender-matching school environments can be a useful policy instrument that promotes female talent in STEM fields, but the positive effect is not universal and thus cannot be generalized for everyone.


Keywords: gender-matching effects; student-teacher gender-matching; single-sex schooling; cognitive performance; non-cognitive performance; education production functions; propensityscore matching; South Korea
JEL-codes: C31; I21; I24; J16; O53

[^0]1. Introduction

Despite noticeable improvement in recent female education and their social status, labor market outcomes show that women are still disadvantaged in earnings and are promoted less frequently to decision-making positions. One explanation for such gender disparity is that men and women behave differently in competitive environments. Women $\hat{\mathscr{Q}}$ underperformance and underrepresentation in competition could explain gender wage differentials and gender gaps in decision-making positions. In an attempt to shed light on the causes of gender-asymmetric behaviors in competition, recent literature has focused on how the gender composition of one $\hat{\Theta}$ surroundings influences individual performance. The empirical evidence suggested in the literature reveals that women tend to underperform when they have to compete with men, and a higher presence of female peers (or the absence of male counterparts) can therefore create a favorable environment for women and girls to boost their performance and competitiveness (Antonovics et al. 2009; Niederle and Vesterlund 2007).

In this respect, single-sex schooling is often considered a policy instrument in education to improve girlsô performance. This is not only because single-sex schools are expected to maximize boysôand girlsôacademic performance by providing them with instructions based on their gender-specific needs, but also because single-sex schooling can strengthen girlsô competitive attitudes. In all-girls schools, girls can be freer from traditional gender roles and can adopt more positive and active gender identities by interacting with female peers who are more present and driven towards lead roles that become more available in the absence of boys (see Bracey 2006; Hill 2015; Gneezy et al. 2003; and Niederle and Vesterlund 2007 for detailed discussions). In literature, various studies provide empirical evidence that single-sex schooling has a positive effect on girlsôstudies and attitudes in different countries: Booth and Nolen (2012a, b) for the United Kingdom; Schneeweis and Zweimuller (2012) for Austria; McCoy et al. (2012) for Ireland; Hoxby (2000), Fryer and Levitt (2010), and Lavy and Schlosser (2011) for the United States; Eisenkopf et al. (2015) for Switzerland; and Jackson (2012) for Trinidad and Tobago. ${ }^{1}$

[^1]However, other studies challenge the idea that single-sex schooling can improve girlsô performance (see Billger 2009; Halpern et al. 2011; Aedin et al. 2013; Strain 2013; and Goodkind et al. 2013). These studies argue that the presumed positive outcome of single-sex schools is likely caused by endogenous school choices, in that students in single-sex schools are systematically different in their characteristics and performance records from students in coeducational schools. This identification issue is taken into account in this paper and accordingly, unobserved heterogeneity that causes self-selection biases is addressed by the following ways. First, an extensive set of educational inputs is incorporated in the empirical model in order to control for individual and school heterogeneity in a holistic manner. Second, a propensity-score matching method is applied to gauge an average treatment effect of attending a single-sex school by equilibrating characteristics of students between single- and mixed-sex schools. Third, the model is estimated by using the sub-sample of students in public schools only in order to minimize differences between schools.

In addition to single-sex schooling, another important type of gender-matching in school environments must be considered: interactions between students and teachers of the same sex. In literature, a growing number of studies have documented that female teachers/professors have a positive influence on female studentsôperformance (Carrell et al. 2010; Nixon and Robinson 1999; Bettinger and Long 2005; Dee 2007). This is likely because they can employ teaching methods and forms of communication that are more suited for girls due to their own gender knowledge. Furthermore, female teachers represent active female professionalism and mentorship for girls and therefore can become positive gender role models for girls, motivating them to actively participate in class and set ambitious career goals. Considering these arguments, this paper investigates single-sex schooling and interactions of a female student and a female teacher as crucial gender-matching school environments that influence girlsôperformance and attitudes. ${ }^{2}$

To identify such gender-matching effects, this paper focuses on South Korea for the empirical analysis. Empirical evidence suggested by the Program for International Student Assessment (PISA, OECD 2015) illustrates an interesting case of gender gaps in studentsôperformance in

[^2]South Korea. Namely, South Korean girls outperform boys in cognitive studies including science, however, they are less motivated and interested than boys in pursuing competitive study and career paths in science, technology, engineering, and math (STEM) ï see Figure 1. This disparity signals that girlsôlow level of motivation and interest in science cannot be explained by their supposed lack of abilities. Rather, it can be attributed to societal environments that discourage women $\hat{s}$ participation in STEM fields. In this paper, STEM fields are at the forefront of the investigative analysis because science and engineering are among the most male-dominated areas that exacerbate gender wage differentials and women $\hat{\propto}$ underrepresentation in competitive labor markets.

By examining the case of South Korea, this paper aims to find whether gender-matching school environments can be instrumental to reduce gender gaps in competitive attitudes in STEM fields. To the present, literature reports a generally positive effect of attending an all-girls school in South Korea, but most analyses focus only on the effect on cognitive outcomes (Park et al. 2011, 2013; Kim 2012; Kim and Law 2012). In addition to single-sex schooling, Lim and Meer (2017) document that female teachers have a positive influence on girlsôstudies in South Korea. On the other hand, Lee et al. (2014) is one of rare studies that investigate the effect of single-sex schools on non-cognitive outcomes in South Korea. Through their analysis of middle-schools in Seoul, they find that girls in all-girls schools are not more competitive than girls in mixed-sex school, and coeducation rather reduces gender gaps in competition. Their findings are contrary to the presumed positive role of single-sex schooling in boosting girlsô confidence and motivation. Given that the effect of single-sex schooling on non-cognitive outcomes is seldom evaluated in South Korea, Lee et al.ôs study is considered a seminal analysis with which findings should be coupled with data of different school levels and wider ranges of regions. Keeping this in mind, my paper builds on evidence of gender-matching school effects on girlsôattitudes by extending the analysis to high school students in different regions throughout South Korea.

Using the 2015 PISA outcomes for South Korea, the findings show that gender-matching school environments can promote girlsômotivation and interest in science. The positive effects of such school environments are largest for girls at the highest tail of science studies. By attending an allgirls school and being taught by a female science teacher, high performing girls become as motivated and interested in pursuing studies and careers in STEM fields as boys. In addition,
female teachers generally have a positive effect on developing competitive attitudes of girls regardless of their study records. However, the effect of single-sex schooling is varying depending girlsôcognitive performance. While maintaining its positive effect on girls at the two ends (high and low) of study performance in science, attending an all-girls school can be detrimental to the attitudinal development of median girls. These findings corroborate that gender-matching school environments can be a useful policy instrument to motivate female talent to pursue careers in STEM fields. However, this positive effect is not universal for all students.

## 2. Empirical Framework

### 2.1. Education Production Function

The central question for the empirical analysis of this paper is to identify the net effect of gendermatching school environments on girlsôperformance. To isolate this effect, the estimation model includes an exhaustive list of covariates that have potentially compounding effects on outcome variables. The selection of variables follows the education production function suggested by Hanushek (1986) and Krueger (1999). In the education production model, outputs (studentsô performance) are determined by the following input:

> Y (educational output) = f (individual; family; school; teacher; and peer inputs)

In this model, studentsô performance $(\mathrm{Y})$ includes not only their study outcomes (cognitive performance), but also attitudinal development (non-cognitive performance) because both cognitive and non-cognitive skills are important determinants of successful career development. The above education production function is rewritten in an econometric model below specifying covariates and their relationships with the outcome variables (see Equation 1):

$$
\begin{align*}
& \text { Performance }_{i}=a+反_{1} \text { female student }_{i}+\text { б }_{2} \text { boy school }{ }_{i}+\text { б }_{3} \text { girl school }_{i}+\text { }_{4} \text { female teacher }{ }_{i} \\
& +W_{5} \text { female studenti }{ }_{i}{ }^{*} \text { female teacher }{ }_{i}+X_{i}{ }^{\prime} \tilde{u}+S_{i}{ }^{\prime} Q+T_{i}{ }^{\prime} 7+B_{i}{ }^{\prime} \emptyset+R_{i}{ }^{\prime} d+u_{i} \tag{1}
\end{align*}
$$

The set of the outcome variables consists of several indicators that evaluate different dimensions of studentsôperformance. First, studentsôcognitive performance is measured by their PISA test
scores in reading, math, and science subjects. Second, the non-cognitive performance of students is proxied by a student $\hat{\propto}$ self-assessment of his/her instrumental motivation, confidence, and interest in science that are available in the PISA survey. These variables reflect important individual attitudes that influence oneô decision to pursue studies and careers in STEM fields:

Cognitive performance $=\{$ science score, math score, reading score $\}$
Non-cognitive performance $=\{$ instrumental motivation, confidence, interest in science $\}$

In estimating the model of non-cognitive performance, a variable that reflects a student $\hat{Q}$ intellectual abilities is additionally included as an explanatory variable because oneô knowledge level leverages his/her non-cognitive performance. Scientific knowledge can be the most important factor of determining attitudes towards science. However, the available measurement of scientific knowledge $\bar{i}$ science (or math) scores in the PISA test $\bar{i}$ has a tautological relationship with non-cognitive performance in science, sharing their latent concepts to a great extent. To avoid this problem, PISA reading scores are used as a proxy to capture a general level of intelligence instead. High correlation between the two scores strengthen the validity of a reading score as a proxy variable: corr. (reading score'science score) $=0.85$. Accordingly, the model of non-cognitive performance takes the following form:

$$
\begin{align*}
& +\mathrm{b}^{\prime}{ }_{4} \text { female teacher }{ }_{i}+\mathrm{b}_{5}^{\prime} \text { female student } t_{i}{ }^{*} \text { female teacher } \mathrm{r}_{\mathrm{i}}+\mathrm{b}^{\prime}{ }_{6} \text { reading score }{ }_{i} \\
& +X_{i}{ }^{\prime} \tilde{u}^{\prime}+S_{i}{ }^{\prime} \mathrm{Q}^{\prime}+\mathrm{T}_{\mathrm{i}}{ }^{\prime}{ }^{\prime}{ }^{\prime}+\mathrm{B}_{\mathrm{i}}{ }^{\prime} \mathrm{D}^{\prime}+\mathrm{R}_{\mathrm{i}}{ }^{\prime} \mathrm{d}^{\prime}+\mathrm{u}^{\prime}{ }_{i}
\end{align*}
$$

Female student is a dummy variable indicating a studentô gender. Female teacher refers to whether student i is taught by a female teacher in the respective course. Boy school and girl school represent single-sex schooling for boys and girls, respectively. Hence, gender-matching effects for girls are estimated through two variables: girl school and female student*female teacher (i.e. a girl is taught by a female teacher). Accordingly, positive gender-matching effects on girlsôperformance are hypothesized as following.

| $\mathrm{H}_{0}:$ | $\boldsymbol{\sigma}_{3( }\left(\boldsymbol{\sigma}_{3}\right)>0$ |
| :--- | :--- |
| $\mathrm{H}_{0}:$ | $\mathrm{\sigma}_{5}\left(\boldsymbol{\sigma}_{5}\right)>0$ |

The model includes an exhaustive set of input variables so that omitted variable biases can be minimized. Accordingly, vectors, $\mathrm{X}, \mathrm{S}, \mathrm{T}, \mathrm{B}$, and R , consist of the following variables: a student $\hat{\bigotimes}$ socioeconomic and family backgrounds (X), school characteristics (S), teachersô characteristics (T), a student $\hat{\varrho}$ behavioral factors (B), and teacher and peer relationship (R). The choice of input variables follows the literature. Studentsôsocioeconomic characteristics are taken from Hanushek (1986) who emphasizes their important role in determining studentsô performance level. The choice of school inputs follows Krueger (2003) and Hanushek (2011) who propose class sizes and teachersô quality as important inputs. In addition, a student $\hat{\bigotimes}$ behavioral patterns and relationships with teachers and peers are incorporated in the model because these variables often mirror a student $\hat{\bigotimes}$ personality and mentalities. Therefore, accounting for such effects can reduce the influence of a student $\hat{\bigotimes}$ unobserved characteristics on his/her performance. The list of input variables in each vector is detailed below and the descriptive statistics of all variables used for the estimations are presented in Appendix A.

- $\quad$ (gender and gender-matching variables) $=$ \{female student; boy school; girl school; female teacher; and female student*female teacher $\}$
- $X$ (student $\hat{\varrho}$ and family characteristics) $=\{$ father $\hat{Q}$ education; mother $\hat{Q}$ education; student $\hat{Q}$ economic, social, and cultural status; family spending on education; parental support for learning at home; parental emotional support; and intellectual ability (in the model of non-cognitive performance) $\}$
- $S$ (school characteristics) $=$ \{public school; community size; student-teacher ratio; school size; and school quality perceived by parents $\}$
- T (teacher $\hat{\alpha}$ characteristics) $=\{$ teacher $\hat{\alpha}$ tenure; and teacher $\hat{\widehat{O}}$ years of teaching experience\}
- $B$ (student $\hat{Q}$ behaviors) $=\{$ frequency of skipping classes; coming to school late; chatting online at school and outside of the school; and participation in social networks $\}$
- R (teacher and peer relationships) $=$ \{degree of feeling belonging to school; and unfairness of teachers $\}$

The econometric models formulated in Equation 1 and 1' are estimated by two methods. First, an OLS estimation is applied, assuming the linearity of the models with continuous dependent variables. Thereafter, the models are constructed as a multilevel (mixed) one in which observations are nested within schools. This approach allows us to account for heterogeneous data patterns across schools. In this model, intercepts are treated as random effects considering the data structure grouped by school. In addition, robust standard errors are clustered at the school level because (unobserved) variations of observations in the same school are possibly correlated to one another.

Estimations are first conducted with the full sample. The full sample includes 3,259 students in 105 high schools in South Korea who took part in PISA in 2015 ( $\mathrm{n}=3,259$; $\mathrm{s}=105$ ). Then, the sample is limited to public schools ( $\mathrm{n}=2,280 ; \mathrm{s}=74$ ). The sample of public schools can reduce unobserved heterogeneity of students and schools because students in private schools tend to be different in their socioeconomic backgrounds from others in public schools, and private schools have different school management systems including own recruitment processes of teachers. ${ }^{3}$

### 2.2.Self-selection into a Single-sex School and Propensity-score Matching

In identifying the net effect of single-sex schooling from other compounding influences, a critical question remains to be answered: are students in single-sex schools systematically different from others in mixed-sex schools? If students decide to attend an all-girls or all-boys school because of their distinguished backgrounds and characteristics, the choice of school is endogenous to their performance outcomes. Under the presence of such self-selection biases, a causal effect of singlesex schooling on studentsô performance cannot be identified. Furthermore, unobserved heterogeneity of schools in terms of curriculum, teachersôquality, and school orientations that may differ between the school types also makes it difficult to isolate the gender-matching effect of single-sex schools.

[^3]In this respect, the data from South Korea provides an advantage of equilibrating students between single- and mixed-sex schools because single-sex schooling is relatively common in the country. In the sample of the PISA 2015, 29.4 percent of boys and 24.8 percent of girls attend single-sex high schools. Thus, systematic differences in students between the two types of schools are less salient in South Korea. Furthermore, a sufficient number of observations of single-sex school attendees can be obtained for a viable comparison. However, the South Korean sample is also not completely free of selection biases because students are not randomly assigned to schools but they can select a preferred one to attend. For instance, since 2010, middle-school students in Seoul have been allowed to submit three names of preferred high schools and they are assigned to a school based on their preferences. According to Kim (2012), students clearly avoid mixed-sex schools after the introduction of this policy. One of the main reasons for preferring single-sex schools is that students in all-boys and all-girls schools outperform others in mixed-sex schools in university entrance exams. Thus, students and parents who are more concerned about their study records and opportunities for higher education are more likely to choose single-sex schools.

With this in mind, a formal analysis is conducted in this section to find whether students in single-sex schools have systematically different characteristics compared to those in mixed-sex schools. To identify a mechanism of attending a single-sex school, various ex-ante conditions of students prior to entering a high school are included in the parsimony and extended models and the effects are estimated by using a probit method (with a binary dependent variable that has a value of 1 if student i attends a single-sex school). The results are presented in Appendix C. Contrary to commonly-held beliefs of single-sex schools, a higher socioeconomic status of a student does not increase the probability of one attending a single-sex school. Further, higher levels of a father $\hat{ß}$ education and home possessions (e.g. TV and wash machines) even reduce a girl̂̂ probability of choosing an all-girls school. Most other variables reflecting studentsô characteristics have no significant effects on single-sex schooling; an exception is family support for children. Educational resources available at home increase a girlô probability to attend an allgirls school, and parental emotional support increases a boyô single-sex schooling chances. In addition, living in a larger city positively affects attendance in an all-boys school ï a logical consequence given that small towns may not be able to host both all-boys and girls schools. Weak relationships between studentsôsocioeconomic conditions and single-sex schooling show
that students are relatively comparable across the two school types in South Korea. Nonetheless, there is some evidence that students whose parents are more supportive are more likely to choose a single-sex school. This finding necessitates the utilization of a method that can empirically address the endogenous school choice. Accordingly, the following approaches are employed in this paper. First, various educational inputs are incorporated in the empirical model in a holistic manner based on the education production function (Hanushek 1986; Krueger 1999) ï see Section 2.1 for details. Second, the sample is limited to public schools only to minimize school heterogeneity. Third, students are matched based on their propensity to attend a single-sex school and an average treatment effect is computed by a propensity-score matching method (PSM).

In the PSM estimations, attending a single-sex school is designated as a treatment effect. An individual $\hat{\Theta}$ probability of choosing single-sex schooling is predicted based on observed covariates, and students with similar probabilities but receiving different treatments (single- or mixed-sex schooling) are matched to equate differences between the treatment and control groups. The average treatment effect (ATE) of attending a single-sex school is then computed by imputing the missing potential outcome for each subject (see Equation $2 / 2^{\prime}$ below). This is done by averaging outcomes of similar subjects that receive the other treatment level. Hence, the PSM estimator captures the average difference between the observed and potential outcomes for each subject (Abadie and Imbens 2011).

$A T E=E[$ non-cognitive outcome single-sex Ï non-cognitive outcome mixed-sex $\mathrm{DG} ; \mathrm{X} ; \mathrm{S} ; \mathrm{T} ; \mathrm{B} ; \mathrm{R}]$
3. Results

### 3.1. Baseline Estimates

In investigating gender gaps in cognitive and non-cognitive performance, descriptive differences between boys and girls are first compared. Overall, South Korean girls do not perform worse than boys in their studies. There is no gender difference in science and math scores measured by mean, median, and distribution differences, while girls outperform boys in reading significantly (see

Appendix B). ${ }^{4}$ In addition to the overall comparisons presented in Appendix B, Figure 2 illustrates quantile distributional plots of performance of boys and girls and shows minimal gender differences in science and math test scores in all fractions of the data.

However, the findings are quite different in non-cognitive performance. Boys express higher levels of instrumental motivation and interest in pursuing careers in STEM fields than girls. ${ }^{5}$ This is evident in all of mean, median, and distribution differences (Appendix B). Also, the quantile analyses depicted in (4) and (6) of Figure 2 reveal that boys have higher levels of motivation and interest in science in all fractions of data points. Such descriptive gender differences in cognitive and non-cognitive performance tentatively suggest that boys are overly motivated and interested in science compared to girls although there is little evidence that they actually outperform girls in this field of study.

In this section, such gender (in)differences found in the descriptive comparisons are further investigated through systematic analyses. In doing so, the following questions are specifically addressed: (i) dose a studentô gender influence his/her cognitive and non-cognitive performance? and (ii) can gender-matching school environments reduce the presumed gender gaps in noncognitive performance?

First, the results of cognitive outcomes estimated with the full sample are presented in Table 1. As expected from the descriptive analysis, a student $\hat{O}$ gender is not a significant determinant of his/her science and math scores. However, in reading, girls have a significant advantage of having a score that is five percentage-points (p.p.) higher than boysôon average. On the contrary, a teacher $\hat{O}$ gender is widely unimportant to studentsô cognitive performance. There is some evidence that being taught by a female science teacher increases both girlsôand boysôtest scores in this subject. However, this result is found in the multilevel estimations only, and the estimated effect is too small to draw an economic meaning. Further, gender-matching environments do not

[^4]influence a studentô cognitive performance. Neither single-sex schooling nor the interaction of a female student and a female teacher has an effect on any of the three test scores.

Among school inputs, a higher student-teacher ratio deteriorates study outcomes, supporting the benefits of small classes. However, a school size (the number of students) is not an important determinant of a studentô study results. School quality evaluated by parents has a positive effect in the OLS estimations but the size of the effect is negligible. A schoolôs status (public or private) and location (community size, indicating urban or rural) are also unimportant. Moreover, teachersô credentials have little to do in improving studentsôstudy performance. Neither teacherô tenure nor years of experience has any effect on test scores. ${ }^{6}$ Limited roles of school and teachersôinputs mirror the importance of private after-school tutoring that often overshadows formal schooling in South Korea (Kim 2012).

In contrast, a studentô family backgrounds and socioeconomic status are important inputs of one $\hat{\beta}$ cognitive performance. A student $\hat{\aleph}$ economic, social, and cultural (ESC) status, family spending on education, and parental emotional support have positive effects on all of science, math, and reading scores. Increasing one-standard deviation in these variables increases the test scores by 3.3 ï 4.8 p.p. (ESC status), $1.2 i ̈ 2.7$ p.p. (family spending), and 0.7 ï 0.9 p.p. (parental emotional support). The effects of one $\hat{\propto}$ ESC status and family spending on education are greater for math scores, while that of parental emotional support has a larger effect on science. In addition, a motherố education positively influences a student $\hat{\AA}$ math score. A one-standard deviation increase in the level of a motherô education increases her childốs math score by 0.7 i 0.9 p.p. Also, evidence suggests a positive effect of a motherô education on her childô science score (but the effect is significant in the OLS model only). On the other hand, a fatherô education does not affect his child $\hat{\propto}$ study outcomes in any of the three subjects. Its effect is possibly surpassed by the effect of a motherô education that tends to share similar features.

Furthermore, a student $\hat{\Theta}$ behaviors are suggested to have great explanatory power over his/her cognitive performance. Frequently skipping classes and coming to school late result in low test

[^5]scores, as does frequenting online chatting at school. However, online activities outside of the school (online chatting out-of-school and social network participation) do not harm a student $\hat{\Theta}$ study performance. In addition, a studentô relationship with teachers is important to the cognitive learning of the student. Increasing distrust in the fairness of a teacher by one standard deviation reduces a studentô test scores by more than 2 points ( 0.2 p.p.) in all three subjects.

So far, the analysis of cognitive performance suggests little support to gender or gender-matching effects. However, the outlooks are different when the effects are estimated on non-cognitive outcomes (see Table 2). First, the gender effect becomes negative for girls. Being a girl, the level of her instrumental motivation in science is lower than boys by $6.5^{\circ} 9.5$ p.p. and that of her interest in science by 6.8 ï 9.2 p.p. On the other hand, a teacher $\hat{O}$ gender has a conflicting effect on boys and girls. Female science teachers reduce boysômotivation, confidence, and interest in science by $4,1.7$, and 1.7 p.p., respectively. However, for girls, the effect of a female teacher is generally positive, as the positive interaction effect of a female student-teacher pair outweighs the negative effect of a female teacher. Specifically, girls increase their motivation and interest in science by 0.7 and 2 p.p., respectively when they are taught by a female science teacher. This positive interaction effect also reduces the negative effect of a girl $\hat{\Theta}$ own gender. If a girl is taught by a female teacher, the negative effect of being a girl on her motivation decreases by 5.3 ï 7.5 percent and that on her interest in science by $20.5{ }^{-1} 21.5$ percent (see Columns 2 and 4 ; and 10 and 12 of Table 2, respectively).

Different from the positive gender matching effect of a female student-teacher pair, single-sex schooling does not influence one $\hat{ß}$ non-cognitive performance ï similar to its minimal role in cognitive outcomes above. Most other school inputs also have no effect, except perceived school quality that is positively associated with oneô motivation and interest levels. Teachersôinputs are also irrelevant to a studentố non-cognitive performance. On the other hand, studentsô family backgrounds and their behavioral patterns provide significant explanations for their attitudinal performance. In particular, a studentô ESC status increases his/her confidence and interest in science, and parental support for learning improves studentsônon-cognitive performance in all three dimensions. However, family spending on education and parentsôemotional support that were important to cognitive outcomes do not play a significant role in enhancing studentsônoncognitive performance. In addition, onê̂ intelligence level (proxied by log reading scores) has a
significant effect on his/her non-cognitive performance, as expected, but the magnitude of the effect is trivial $i$ ï about a tenth of 1 p.p.

Among behavioral patterns, frequently coming to school late and participating in social networks negatively influence non-cognitive performance in all three dimensions. Increasing frequency in late coming by one-standard deviation decreases a studentố non-cognitive outcomes by 0.9 ï 1.1 p.p. Frequenting in social networks by the same margin reduces one $\hat{\alpha}$ motivation, confidence, and interest levels by $2.34,0.8$, and 1.74 p.p., respectively. Online social networks are negatively associated with studentsôattitudes possibly because some students may use SNS as a tool to shy away from activities in the real world. In contrast, chatting online at school reinforces a student $\hat{\Theta}$ positive attitudes. Increasing one-standard deviation in the frequency of in-school online chatting improves one $\hat{Q}$ non-cognitive performance by 1.1 IÏ 1.2 p.p. One plausible interpretation on this result is that online chatting inside the school mirrors a studentô outgoing involvement in school activities. Examining detailed mechanisms of online activities and their effects on one $\hat{\not}$ attitudes is beyond the scope of this paper. However, given the rising importance of online behaviors on cognitive and non-cognitive development of the youth today, this finding certainly requires further research. In addition to online behaviors, feeling belonging to school increases all three dimensions of non-cognitive outcomes, while the unfairness of a science teacher decreases one $\hat{\Theta}$ interest in this subject.

### 3.2.Effects of Unobserved Heterogeneity

The baseline estimations in Section 3.1 utilize a large number of control variables in order to single-out net gender-matching effects. Among the two gender-matching indicators, the interaction of a female student-teacher pair is assumed to be fairly exogenous because the assignment of teachers inside a school is a decision of the school but not of students/parents. One may speculate that female teachers may be assigned to systematically different classes ï for example, consisting of worse-performing students or those with low-income family backgrounds. However, this is less likely. In South Korea, students are randomly allocated among different classes (at least in regular classes that are surveyed in the PISA), independent of their performance or characteristics. Thus, each class includes wide ranges of students of different study ranks and backgrounds. On the other hand, single-sex schooling is more likely endogenous
to studentsôperformance given that students/parents can select a list of preferred schools and are assigned to one of them. Including an extensive set of covariates helps reduce biases arising from an endogenous school choice. However, a large set of controls does not fully ensure that no covariate remains unobserved. For instance, unobserved family values and studentsôpersonality may affect their performance and school choice simultaneously. Hence, the endogenous choice of single-sex schooling (self-selection) is further examined by employing the following approaches in this section.

First, the sample is limited to public schools. Students in private schools are likely different from those in public schools because the former tend to come from more selected family backgrounds. Furthermore, private schools have more heterogeneous school quality and curricula decided by their founders and management instead of the state. Thus, limiting the sample to public schools can reduce unobserved heterogeneity of students and schools that influences studentsô performance. The public-school sample represents 70 percent of the full sample.

The results of cognitive outcomes in public schools are presented in Table 3. The results are similar to those of the full sample in Table 1. Most importantly, the gender-matching school environments play no role in any of the three subjects in public schools. Further, there is no negative gender effect on cognitive performance. Neither being a female student nor being taught by a female teacher reduces a studentô study outcomes. Instead, being a girl has a positive effect on her reading score. The effects of school, family and behavioral inputs remain similar to those of the full sample. However, perceived school quality loses its significant effect in public schools. This finding implies that public schools offer more homogeneous quality of education than private ones.

On the other hand, the results of non-cognitive performance in public schools have three noticeable differences compared to the full-sample (see Table 4). First, single-sex schooling increases girlsô motivation in science in public schools ï different from no effect in the full sample. Second, a female science teacher has no longer a negative effect on boysôinterest in science. Third, the positive effect of a female science teacher is twice as large for girls in public schools as in the full sample. These differences suggest that the gender-matching effects are more important in ordinary environments of public schools than arguably more privileged private ones.

Furthermore, while the effect of a girl̂$\hat{\widehat{O}}$ own gender remains negative in public schools, the positive gender-matching effects with larger magnitudes significantly reduce the negative gender effect. For example, if a girl attends an all-girls school and is taught by a female science teacher, the negative effect of her gender on her motivation in science decreases by $70 і ̈ \quad 80$ percent (Columns 2 and 4). Also, the positive interaction effect of a female student-teacher pair reduces the negative gender effect on a girl̂̂ interest in science by 40 percent (Columns 10 and 12).

As a second method of addressing the effects of unobserved heterogeneity, propensity-score matching is used, as discussed in Section 2.2. Table 5 presents the PSM results, in that the effect of an all-girls school is generally minimal in determining girlsôperformance. In the full sample, the only significant effect arises in girlsômath scores but the effect is negative and marginally significant at a 10 percent level only. Further, this effect does not hold in the public school sample. In public schools, the positive effect of attending all-girls schools on girlsômotivation that was found in Table 4 loses its significance in the PSM analysis, while the effect becomes positive for girlsôconfidence in this estimation (but it is significant at a 10 percent level only).

As seen here, the PSM analysis provides little support to the role of all-girls schools, but attending an all-boys school has a more significant effect on boysôcognitive performance. The average treatment effect is positive for boys, increasing their science and reading scores by 3.5 and 3.9 p.p., respectively (Table 5.1). However, when the sample is limited to public schools, the positive effect of all-boys schools disappears. The positive effect is, indeed, driven by those who selected private boy schools, but not by the general population of boy students in South Korea.

### 3.3.Response Heterogeneity by Cognitive Performance

The results above provide some but not very robust evidence for gender-matching effects on girlsô performance. The gender-matching of a female student-teacher pair maintains its positive effect on girlsô motivation and interest in science. However, the effect of all-girls schools is not consistent, particularly losing its significance after taking into account self-selection effects.

While the analysis has so far focused on the gender-matching effects at the aggregate level, the effects may be heterogeneous depending on studentsôstudy performance. Specifically, girls with
higher cognitive abilities would benefit more from gender-matching environments because in such surroundings, girlsôabilities are likely fairly valued and thus high performing girls would be encouraged to be confident and aim high. This argument is plausible because the absence of male counterparts enables girls and women to be freer of their traditionally defined gender role of being modest, and therefore they can develop more competitive attitudes (Niederle and Vesterlund 2007). With this mind, the gender-matching effects are further investigated in this section by disentangling the effects based on studentsô study outcomes. This approach lends insights into potentially heterogeneous responses to single-sex schooling and the interaction of a female student and a female teacher.

To estimate heterogeneous gender-matching effects, students are sub-grouped based on the quartiles of their science scores: the $4^{\text {th }}$ quartile (science score $>582$ ); the $3^{\text {rd }}$ quartile ( $518<$ science score < 582); the $2^{\text {nd }}$ quartile ( 449 < science score < 518) ; and the $1^{\text {st }}$ quartile (science score < 449). Table 6 shows the results of the sub-sample estimations including both public and private schools. ${ }^{7}$ Consistent with the results at the aggregate level in Table 2, a girl̂̂́s own gender constrains her from being motivated and interested in science regardless of her science score. Notably, the negative gender effect is largest among the best performing girls (the $4^{\text {th }}$ quartile). For instance, the negative effect of being a girl on a girlô motivation is twice as large for the best performing girls as the worst performing ones (the $1^{\text {st }}$ quartile). Also, it is 20 percent larger on the interest of the best performing girls than that of the worst ones. This finding that a girlô gender has the most detrimental influence on the best performing girls implies that high performing women are discredited in society instead of being recognized (Cho 2017).

However, the negative gender effect on a girl can be mitigated through the gender-matching effects. Furthermore, the positive gender-matching effects are largest among the best performing girls in science. For instance, if a girl in the $4^{\text {th }}$ quartile is taught by a female teacher, the negative gender effect decreases by 26 percent for her instrumental motivation, and by nearly 50 percent for her interest in science (Columns 1 and 9). Also, there is evidence of the positive interaction effect of a female student-teacher pair on girls in the $1^{\text {st }}$ and $2^{\text {nd }}$ quartiles. Being taught by a

[^6]female teacher suppresses the negative gender effect on girls in the $1^{\text {st }}$ quartile and they become more motivated than boys in the same group (Column 4). Similarly, the positive interaction effect reduces the negative gender effect on the interest of girls in the $2^{\text {nd }}$ quartile by 80 percent (Column 11). However, there is no such gender-matching effect found in girls in the $3^{\text {rd }}$ quartile.

In addition, single-sex schooling further moderates the negative gender effect on girls $\ddot{i}$ particularly the best performing ones. Heterogeneous responses to single-sex schooling are identified by a propensity-score matching method (see Table 8.1). The effect is positive for girls in the $4^{\text {th }}$ quartile and this effect maintains in all three dimensions of non-cognitive performance. The magnitudes of the effect are considerably large ( $0.154 \mathbf{I}^{0.443}$ ), thus the negative effect of a girl̂̂ own gender can be minimized to a great extent (by 50 ï 90 percent). Moreover, combining both gender-matching effects of single-sex schooling and female student-teacherô interactions, the best performing girls can be more motivated than the best performing boys, and also they can be (almost) as interested in science as boys in the same group. For example, by attending an allgirls school and being taught by a female teacher, the net gender effect of a girl in the $4^{\text {th }}$ quartile on her instrumental motivation become $+0.079,{ }^{8}$ in that her motivation level is about 2.2 p.p. higher than a boy in the same quartile. Also, the negative gender effect on her interest in science is nearly nullified in such gender-matching environments.

Moreover, single-sex schooling has a positive effect on the motivation and confidence of the worst performing girls. The negative gender effect on the motivation of a girl in the $1^{\text {st }}$ quartile is almost offset by the positive effect of single-sex schooling. Further, if she is taught by a female teacher in an all-girls school, her motivation level is $6.5 \mathrm{p} . \mathrm{p}$. higher than boys in the same quartile. Also, by attending an all-girls school, the confidence level of the worst performing girls becomes 3.4 p.p. higher than the worst performing boys. But, the gender-matching environments do not influence the interest level of the worst performing girls, different from the positive effect on the best performing ones.

On the other hand, for median girls who are ranked in the $3^{\text {rd }}$ and $2^{\text {nd }}$ quartiles in their science scores, single-sex schooling has an opposite effect. Attending an all-girls school negatively

[^7]affects the confidence and interest levels of girls in the $3^{\text {rd }}$ quartile. While there is generally no difference in confidence between boys and girls, girls in this quartile who attend an all-girls school have a confidence level that is 10 p.p. lower than that of boys and girls in the same quartile who are in coeducational schools. Also, the negative gender effect on girlsôinterest in science becomes more than twice larger if a girl in this quartile attends an all-girls school ï i.e. from ï 4.5 to $i ̈ 10.8$ p.p. For girls in the $2^{\text {nd }}$ quartile, single-sex schooling is less detrimental than others in the $3^{\text {rd }}$ quartile, but it still reduces their confidence level by 5 p.p.

On the boysôside, the effect of single-sex schooling is insignificant to a large extent. However, there is evidence that attending an all-boys school decreases the confidence level of better performing boys who are ranked in the $4^{\text {th }}$ and $3^{\text {rd }}$ quartiles. These results suggest that gendermatching schooling is more beneficial to the attitudinal development of girls ï particularly the best and worst performing ones, but such benefits accompany costs on median girls and better performing boys who are rather disadvantaged by single-sex schooling.

The heterogeneous responses to gender-matching school environments are further evidenced in the public school sample (Tables 7 and 8.2). The findings in public schools are similar to those of the full sample to a large degree, but there are several differences. First, while the positive effect of a female student-teacher pair generally remains consistent, it loses significance on the interest level of the best performing girls. Second, the positive effect of all girls-schools on the best performing girls is consistent in public schools, but the effect is no longer significant in the worst performing group. Third, single-sex public schooling has no negative effect on median girlsô confidence (in the $3^{\text {rd }}$ and $2^{\text {nd }}$ quartiles), but maintains its negative influence on the interest level of girls in the $3^{\text {rd }}$ quartile. By and large, the evidence in public schools corroborates the heterogeneous gender-matching effects that are particularly more beneficial to the best performing girls.

## 4. Conclusion

If gender-matching environments can create a positive effect on female performance and attitudes, that can be used as a policy instrument to reduce gender inequality in society. My analysis, however, posits complex implications to this question. The effect of gender-matching school
environments is not universal but produces heterogeneous outcomes across different student groups. High-performing girls are the largest beneficiaries of single-sex schooling and female teachers, in that their motivation, confidence, and interest in science are promoted. However, such gains accompany costs on median girls who do not benefit in these environments. Particularly, attending an all-girls school may even discourage median girlsô confidence and interest in science.

The mixed outcomes of single-sex schooling complicate policy-making. If a policy priority is given to promote female talent in STEM fields, gender-matching school environments can be a viable policy option. However, for the purpose of general education that should leave no one behind, single-sex schooling may not be the best choice. In this respect, one may consider the promotion of female science teachers specifically for girls as an alternative policy focus, as the effect of female teachers is positive or at least neutral for all girls. This finding emphasizes the importance of female mentors who can play a positive gender role in motivating girls.

## References

Abadie, Alberto and Guido W. Imbens. 2011. Bias-corrected Matching Estimators for Average Treatment Effects. Journal of Business and Economic Statistics 29: 1-11.

Aedin, Doris, Donal OồVeill and Olive Sweetman. 2013. Gender, Single-sex Schooling and Maths Achievement. Economics of Education Review 35: 104-119.

Akerlof, George and Rachel Kranton 2000. Economics and Identity. Quarterly Journal of Economics 115: 715-753.

Antonovics, Kate, Peter Arcidiacono, and Randall Walsh. 2009. The Effects of Gender Interactions in the Lab and in the Field. Review of Economics and Statistics 91: 152-162.
Bettinger, Eric and Bridget Terry Long. 2005. Do Faculty Serve as Role Models? The Impact of Instructor Gender on Female Students. American Economic Review: Papers and Proceedings 95: 152-157.

Billger, Sherrilyn. 2009. On Reconstructing School Segregation: The Efficacy and Equity of Single-Sex Schooling. Economics of Education Review 28: 393-402.

Booth, Alison and Patrick Nolen. 2012(a). Choosing to Compete: How Different are Girls and Boys? Journal of Economic Behaviour and Organization 81: 542-555.

Booth, Alison and Patrick Nolen. 2012(b). Gender Differences in Risk Behaviour: Does Nurture Matter? Economic Journal 122: 56-78.

Bracey, Gerald. 2006. Separate but Superior? A Review of Issues and Data Bearing on Single-sex Education. Tempe: Educational Policy Research Unit.
Carrell, Scott, Marianne Page, and James West. 2010. Sex and Science: How Professor Gender Perpetuates the Gender Gap. Quarterly Journal of Economics 1101ï 1144.
Cho, Seo-Young. 2017. Explaining Gender Differences in Confidence and Overconfidence in Math. 2017. MAGKS DP No. 01-2017.

Dee, Thomas. 2007. Teachers and the Gender Gaps in Student Achievement. Journal of Human Resources 42: 528-554.

Eisenkopf, Gerald, Zohal Hessami, Urs Fischbacher, and Heinrich W. Ursprung. 2015. Academic Performance and Single-sex Schooling: Evidence from a Natural Experiment in Switzerland. Journal of Economic Behavior and Organization 115: 123-143.

Fryer, Roland G. Jr. and Steven D. Levitt. 2010. An Empirical Analysis of the Gender Gap in Mathematics . American Economic Journal: Applied Economics 2(2): 210-240.

Goodkind, Sara, Lisa Schelbe, Andrea A. Joseph, Daphne E. Beers, Stephanie L. Pinsky. 2013. Children and Youth Services Review 35: 1174-1181.

Gneezy, Uri, Muriel Niederle and Aldo Rustichini. 2003. Performance in Competitive Environments: Gender Differences. Quarterly Journal of Economics 118(3): 1049-1073.

Halpern, Diane, Lise Eliot, Rebecca Bigler, Richard Fabes, Laura Hanish, Janet Hyde, Lynn Liben, and Carol Martin. 2011. The Pseudoscience of Single-sex Schooling. Science 333: 1706-1707.

Hanushek, Eric. 1986. The Economics of Schooling: Production and Efficiency in Public Schools. Journal of Economic Literature 24:1141-1177.

Hanushek, Eric. 2011. The Economic Value of Higher Teacher Quality. Economics of Education Review 30: 466-479.

Hill, Andrew. 2015. The Girl Next Door: The Effect of Opposite Gender Friends on High School Achievement. American Economic Journal: Applied Economics 7(3): 147-177.

Hoxby, Caroline. 2000. Peer Effects in the Classroom: Learning from Gender and Race Variation. National Bureau of Economic Research Working Paper No. 7867.

Huguet, Pascal and Isabelle Regner. 2007. Stereotype Threat among School Girls in QuasiOrdinary Classroom Circumstances. Journal of Educational Psychology 99: 545-560.

Jackson, C. Kirabo. 2012. Single-sex Schools, Student Achievement, and Course Selection: Evidence from Rule-based Student Assignments in Trinidad and Tobago. Journal of Public Economics 96: 173-187.

Kim, Hisam. 2012. Estimation of Education Production Function Focusing on School Effects and Policy Directions for Primary and Secondary Education (written in Korean). KDI Research Monograph 2012-09 (Korea Development Institute): Seoul.

Kim, Doo Hwan and Helen Law. 2012. Gender Gap in Maths Test Scores in South Korea and Hong Kong: Role of Family Background and Single-sex Schooling. International Journal of Education Development 32: 92-103.
Korea Education Development Institute. 2016. Brief Statistics on Korean Education (written in Korean). Ministry of Education of the Republic of Korea: Seoul.
Krueger, Alan. 1999. Experimental Estimates of Education Production Functions. Quarterly Journal of Economics 114: 497-532.

Krueger, Alan. 2003. Economic Considerations and Class Size. Economic Journal 113:34-63.
Lavy, Victor and Analía Schlosser. 2011. Mechanisms and Impacts of Gender Peer Effects at School. American Economic Journal: Applied Economics 3 (2): 1-33.

Lee, Soohyung, Muriel Niederle and Namwook Kang. 2014. Do Single-sex Schools Make Girls More Competitive? Economics Letters 124: 474-477.

Lim, Jaegeum and Jonathan Meer. 2017. The Impact of Teacher-Student Gender Matches: Random Assignment Evidence from South Korea. Journal of Human Resources 52: 979997.

McCoy, Selina, Emer Smyth and Joanne Banks. 2012. The Primary Classroom: Insights from the 'Growing Up in Ireland' Study. Economics and Social Research Institute Working Paper: Dublin.

Niederle, Muriel and Lise Vesterlund. 2007. Do Women Shy Away from Competition? Do Men Compete Too Much? Quarterly Journal of Economics 122(3): 1065ї 1101.
Nixon, Lucia and Michael D. Robinson. 1999. The Educational Attainment of Young Women: Role Model Effects of Female High School Faculty. Demography 36: 185-194.
Organization for Economic Cooperation and Development (OECD). 2015. Program for International Student Assessment (PISA) 2014. OECD: Paris.

Park, Hyunjoon, Jere Behrman and Jaesung Choi. 2011. Causal Effects of Single-Sex Schools on Studentsô STEM Outcomes by Gender and Parental SES. University of Pennsylvania Working Paper.

Park, Hyunjoon, Jere Behrman and Jaesung Choi. 2013. Causal Effects of Single-Sex Schools on College Entrance Exams and College Attendance: Random Assignment in Seoul High Schools. Demography 50: 447-469.
Schneeweis, Nicole and Martina Zweimuller. 2012. Girls, Girls, Girls: Gender Composition and Female School Choice. Economics of Education Review 31: 482-500.
Strain, Michael. 2013. Single-sex Classes \& Student Outcomes: Evidence from North Carolina. Economics of Education Review 36: 73-87.

Table 1. Gender and Gender-Matching Effects on Cognitive Performance, full sample

| Dependent Variable | Log Science Score |  |  |  | Log Math Score |  |  |  | Log Reading Score |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Method | OLS |  | Multilevel |  | OLS |  | Multilevel |  | OLS |  | Multilevel |  |
|  | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) | (9) | (10) | (11) | (12) |
| Female Student | -0.0001 | 0.001 | -0.005 | 0.001 | -0.003 | -0.015 | -0.01 | -0.018** | $0.053^{* * *}$ | 0.048*** | $0.048^{* * *}$ | 0.043*** |
|  | (0.010) | (0.013) | (0.008) | (0.010) | (0.010) | (0.013) | (0.006) | (0.008) | (0.011) | (0.014) | (0.007) | (0.009) |
| Boy School | 0.022 | 0.023 | 0.022 | 0.023 | 0.016 | 0.014 | 0.016 | 0.015 | 0.017 | 0.016 | 0.016 | 0.015 |
|  | (0.021) | (0.021) | (0.023) | (0.023) | (0.020) | (0.020) | (0.022) | (0.021) | (0.022) | (0.022) | (0.024) | (0.023) |
| Girl School | 0.001 | 0.001 | 0.003 | 0.003 | -0.027 | -0.026 | -0.022 | -0.022 | 0.011 | 0.011 | 0.013 | 0.013 |
|  | (0.015) | (0.015) | (0.018) | (0.018) | (0.017) | (0.017) | (0.019) | (0.019) | (0.014) | (0.014) | (0.016) | (0.016) |
| Public School | -0.017 | -0.017 | -0.025 | -0.025 | -0.020 | -0.021 | -0.026 | -0.026 | -0.023 | -0.024 | -0.031 | -0.031 |
|  | (0.016) | (0.016) | (0.019) | (0.019) | (0.017) | (0.017) | (0.020) | (0.020) | (0.019) | (0.018) | (0.021) | (0.021) |
| Community Size | -0.011 | -0.011 | -0.011 | -0.011 | -0.010 | -0.010 | -0.010 | -0.010 | -0.010 | -0.010 | -0.010 | -0.009 |
|  | (0.009) | (0.009) | (0.011) | (0.011) | (0.009) | (0.009) | (0.011) | (0.011) | (0.010) | (0.010) | (0.011) | (0.011) |
| Student-Teacher Ratio | $-1.54 * *$ | -1.53** | -2.02 *** | $-2.01 * * *$ | -1.55** | -1.58** | -2.12 *** | -2.14*** | -1.22* | -1.23* | -1.79** | -1.80 ** |
|  | (0.690) | (0.684) | (0.774) | (0.777) | (0.701) | (0.698) | (0.80) | (0.795) | (0.731) | (0.732) | (0.833) | (0.835) |
| School Size | 0.00001 | 0.00001 | 0.00002 | 0.00002 | 0.00003 | 0.00003 | 0.00004 | 0.00003 | 0.00002 | 0.00002 | 0.00003 | 0.00003 |
|  | (0.00003) | (0.00003) | (0.00003) | (0.00003) | (0.00003) | (0.00003) | (0.00003) | (0.00003) | (0.00003) | (0.00003) | (0.00003) | (0.00003) |
| Perceived School Quality | 0.011** | 0.011** | 0.001 | 0.001 | 0.011*** | 0.011*** | 0.0007 | 0.0007 | 0.011** | 0.011** | 0.001 | 0.001 |
|  | (0.004) | (0.004) | (0.004) | (0.004) | (0.004) | (0.004) | (0.004) | (0.004) | (0.005) | (0.005) | (0.004) | 80.004) |
| Female Teacher | 0.006 | 0.007 | 0.013** | 0.018** | 0.002 | -0.008 | 0.005 | -0.002 | -0.001 | -0.006 | 0.001 | -0.003 |
|  | (0.008) | (0.012) | (0.006) | (0.009) | (0.009) | (0.012) | (0.007) | (0.009) | (0.009) | (0.012) | (0.007) | (0.009) |
| Female Student |  | -0.002 |  | -0.011 |  | 0.020 |  | 0.014 |  | 0.009 |  | 0.008 |
| *Female Teacher |  | (0.014) |  | (0.011) |  | (0.014) |  | (0.012) |  | (0.014) |  | (0.012) |
| Teacherôs Tenure | -0.008 | -0.008 | -0.007 | -0.007 | -0.005 | -0.005 | -0.005 | -0.005 | -0.019 | -0.019 | -0.015 | -0.015 |


|  | (0.011) | (0.011) | (0.009) | (0.009) | (0.013) | (0.013) | (0.009) | (0.009) | (0.013) | (0.013) | (0.009) | (0.009) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Experience of Teacher | -0.0001 | -0.0001 | 0.0004 | 0.0004 | -0.0005 | -0.001 | 0.00006 | 0.00005 | -0.0005 | -0.0006 | -0.0002 | -0.0002 |
|  | (0.0005) | (0.0005) | (0.0004) | (0.0004) | (0.0005) | (0.001) | (0.0004) | (0.0004) | (0.0005) | (0.0005) | (0.0004) | (0.0004) |
| Fatherôs Education | 0.0003 | 0.0003 | -0.004 | -0.004 | 0.004 | 0.003 | -0.0006 | -0.0006 | 0.0006 | 0.0006 | -0.004 | -0.004 |
|  | (0.005) | (0.005) | (0.005) | (0.005) | (0.005) | (0.005) | (0.005) | (0.005) | (0.005) | (0.005) | (0.005) | (0.005) |
| Motherô Education | 0.008* | 0.008* | 0.007 | 0.007 | 0.008** | 0.009** | 0.007* | 0.007* | 0.007 | 0.007 | 0.006 | 0.006 |
|  | (0.005) | (0.005) | (0.004) | (0.004) | (0.004) | (0.004) | (0.004) | (0.004) | (0.005) | (0.004) | (0.004) | (0.004) |
| Economic, Social and | 0.061*** | $0.061^{* * *}$ | 0.043*** | 0.043*** | 0.069*** | $0.069 * * *$ | 0.052*** | 0.052*** | 0.067*** | $0.067 * * *$ | 0.048*** | 0.048*** |
| Cultural Status | (0.008) | (0.008) | (0.007) | (0.007) | (0.008) | (0.007) | (0.007) | (0.007) | (0.008) | (0.008) | (0.007) | (0.007) |
| Family Spending | 0.016*** | 0.016*** | 0.009*** | 0.009*** | 0.020*** | 0.020*** | 0.013*** | 0.013*** | 0.019*** | 0.019*** | 0.012*** | 0.012*** |
| on Education | (0.003) | (0.003) | (0.002) | (0.002) | (0.003) | (0.003) | (0.003) | (0.003) | (0.003) | (0.003) | (0.002) | (0.002) |
| Parental Support for | 0.007* | 0.007* | 0.005 | 0.005 | 0.002 | 0.001 | -0.0004 | -0.0005 | 0.006* | 0.006* | 0.004 | 0.004 |
| Learning at Home | (0.004) | (0.004) | (0.004) | (0.004) | (0.003) | (0.003) | (0.003) | (0.003) | (0.004) | (0.004) | (0.003) | (0.003) |
| Parental Emotional Support | 0.008** | 0.008** | 0.007** | 0.007** | 0.007** | 0.007** | 0.006* | 0.006* | 0.007** | 0.007** | 0.006* | 0.006* |
|  | (0.003) | (0.003) | (0.003) | (0.003) | (0.003) | (0.003) | (0.003) | (0.003) | (0.003) | (0.003) | (0.003) | (0.003) |
| Skipping (some) Classes | $-0.068^{* * *}$ | $-0.068 * * *$ | -0.051** | -0.051 ** | $-0.062 * * *$ | $-0.062 * * *$ | -0.043** | -0.043** | -0.087*** | $-0.087 * * *$ | -0.070 *** | $-0.070^{* * *}$ |
|  | (0.025) | (0.025) | (0.024) | (0.024) | (0.022) | (0.022) | (0.020) | (0.020) | (0.027) | (0.027) | (0.026) | (0.026) |
| Coming to School Late | $-0.027^{* * *}$ | $-0.027 * * *$ | $-0.023 * * *$ | $-0.023 * * *$ | $-0.032 * * *$ | $-0.032 * * *$ | $-0.028 * * *$ | -0.028*** | -0.026*** | $-0.026 * * *$ | $-0.022^{* * *}$ | $-0.022^{* * *}$ |
|  | (0.006) | (0.006) | (0.006) | (0.006) | (0.006) | (0.006) | (0.006) | (0.006) | (0.006) | (0.006) | (0.006) | (0.006) |
| Chatting Online | -0.001 | -0.001 | 0.001 | 0.001 | -0.0007 | -0.001 | -0.001 | 0.001 | -0.002 | -0.002 | 0.00004 | 0.00005 |
| (Outside of School) | (0.002) | (0.002) | (0.002) | (0.002) | (0.002) | (0.002) | (0.002) | (0.002) | (0.002) | (0.002) | (0.002) | (0.002) |
| Participation | -0.004* | -0.004* | -0.003 | -0.003 | -0.002 | -0.002 | -0.001 | -0.001 | -0.002 | -0.002 | -0.001 | -0.001 |
| in Social Networks | (0.002) | (0.002) | (0.002) | (0.002) | (0.002) | (0.002) | (0.002) | (0.002) | (0.002) | (0.002) | (0.002) | (0.002) |
| Chatting Online (School) | $-0.027^{* * *}$ | $-0.027 * * *$ | -0.026*** | -0.026*** | $-0.023 * * *$ | $-0.023 * * *$ | $-0.022 * * *$ | -0.022*** | $-0.029 * * *$ | -0.030 *** | $-0.030^{* * *}$ | $-0.030^{* * *}$ |
|  | (0.005) | (0.005) | (0.003) | (0.004) | (0.005) | (0.005) | (0.004) | (0.004) | (0.006) | (0.006) | (0.004) | (0.004) |
| Feeling Belonging to School | 0.001 | 0.001 | -0.003 | -0.003 | 0.012*** | 0.012*** | 0.009*** | 0.009*** | 0.0003 | 0.003 | -0.003 | -0.003 |


|  |  |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Unfairness of Teacher | -0.008*** | -0.008*** | -0.007*** | -0.007*** | -0.007*** | $-0.007 * * *$ | -0.007*** | -0.007*** | -0.008*** | -0.008*** | -0.007*** | -0.007*** |
|  | (0.001) | (0.001) | (0.001) | (0.001) | (0.001) | (0.001) | (0.001) | (0.001) | (0.001) | (0.001) | (0.001) | (0.001) |
| Number of Observations | 3,259 | 3,259 | 3,259 | 3,259 | 3,258 | 3,258 | 3,258 | 3,258 | 3,258 | 3,258 | 3,258 | 3,258 |
| Number of Schools | 105 | 105 | 105 | 105 | 105 | 105 | 105 | 105 | 105 | 105 | 105 | 105 |
| $\mathrm{R}^{2}$ | 0.246 | 0.246 |  |  | 0.281 | 0.281 |  |  | 0.277 | 0.277 |  |  |
| Wald Chi2 |  |  | 462.3*** | 469.3*** |  |  | 631.8*** | 629.8*** |  |  | 507.7*** | 511.4*** |

Note: Robust standard errors are in parentheses. Robust standard errors are clustered at the school level in the linear estimations. $* p<.10, * * p<.05, * * * p<.001$.

Table 2. Gender and Gender-Matching Effects on Non-cognitive Performance, full sample

| Dependent Variable |  | umental Mo | ation in Scie |  |  | Confiden | Science |  |  | Interest | Science |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Method | OLS |  | Multilevel |  | OLS |  | Multilevel |  | OLS |  | Multilevel |  |
|  | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) | (9) | (10) | (11) | (12) |
| Female Student | $-0.244 * * *$ | -0.349*** | $-0.221^{* * *}$ | $-0.306 * * *$ | 0.038 | -0.007 | 0.038 | -0.005 | $-0.355^{* * *}$ | $-0.465 * * *$ | $-0.351^{* * *}$ | -0.454*** |
|  | (0.046) | (0.055) | (0.044) | (0.053) | (0.046) | (0.073) | (0.045) | (0.073) | (0.038) | (0.055) | (0.036) | (0.054) |
| Boy School | 0.017 | -0.004 | 0.030 | 0.011 | -0.017 | -0.026 | -0.017 | -0.026 | 0.021 | -0.001 | 0.028 | 0.006 |
|  | (0.058) | (0.057) | (0.058) | (0.059) | (0.061) | (0.063) | (0.061) | (0.063) | (0.057) | (0.058) | (0.057) | (0.058) |
| Girl School | 0.091 | 0.093 | 0.079 | 0.080 | -0.095 | -0.094 | -0.094 | -0.094 | 0.056 | 0.057 | 0.055 | 0.057 |
|  | (0.076) | (0.072) | (0.075) | (0.071) | (0.079) | (0.078) | (0.079) | (0.078) | (0.057) | (0.053) | (0.056) | (0.052) |
| Public School | 0.039 | 0.040 | 0.027 | 0.030 | 0.033 | 0.034 | 0.032 | 0.033 | 0.027 | 0.029 | 0.021 | 0.024 |
|  | (0.067) | (0.066) | (0.067) | (0.066) | (0.052) | (0.051) | (0.051) | (0.051) | (0.049) | (0.048) | (0.048) | (0.048) |
| Community Size | 0.018 | 0.022 | 0.017 | 0.020 | -0.008 | -0.007 | -0.008 | -0.007 | 0.005 | 0.008 | 0.006 | 0.009 |
|  | (0.035) | (0.034) | (0.035) | (0.034) | (0.029) | 80.029) | (0.029) | (0.029) | (0.024) | (0.023) | (0.024) | (0.023) |
| Student-Teacher Ratio | -2.037 | -2.234 | -1.932 | -2.096 | -1.128 | -1.213 | -1.157 | -1.231 | -1.637 | -1.854 | -1.699 | -1.899 |
|  | (2.156) | (2.117) | (2.113) | (2.074) | (1.879) | (1.869) | (1.882) | (1.869) | (1.797) | (1.765) | (1.818) | (1.776) |
| School Size | -0.0001 | -0.0001 | -0.0001 | -0.0001 | -0.00001 | -0.00001 | -0.00001 | -0.00001 | -0.00007 | -0.00008 | -0.00008 | -0.00008 |
|  | (0.0001) | (0.0001) | (0.0001) | (0.0001) | (0.00009) | (0.00009) | (0.00009) | (0.00009) | (0.00009) | (0.00009) | (0.00009) | (0.00009) |
| Perceived School Quality | 0.072*** | 0.071*** | 0.070*** | 0.069*** | 0.037 | 0.037 | 0.037 | 0.037 | 0.070*** | 0.070*** | 0.071*** | 0.071*** |
|  | (0.021) | (0.021) | (0.022) | (0.021) | (0.027) | (0.027) | (0.027) | (0.027) | (0.022) | (0.022) | (0.023) | (0.022) |
| Female Teacher | -0.071* | -0.158*** | -0.059 | -0.131*** | -0.083* | -0.119** | -0.082* | -0.118** | -0.005 | -0.094* | 0.003 | -0.084* |
|  | (0.040) | (0.053) | (0.038) | (0.048) | (0.044) | (0.059) | (0.043) | (0.059) | (0.037) | (0.053) | (0.035) | (0.051) |
| Female Student*Female Teacher |  | 0.184*** |  | 0.146** |  | 0.078 |  | 0.075 |  | 0.190*** |  | 0.177*** |
|  |  | (0.070) |  | (0.066) |  | (0.088) |  | (0.088) |  | (0.068) |  | (0.067) |
| Teacherô Tenure | 0.013 | 0.010 | -0.023 | -0.025 | 0.032 | 0.030 | 0.034 | 0.032 | -0.056 | -0.059 | -0.053 | -0.056 |


|  | (0.057) | (0.058) | (0.057) | (0.058) | (0.057) | (0.057) | (0.057) | (0.057) | (0.045) | (0.045) | (0.046) | (0.046) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Experience of Teacher | -0.003 | -0.003 | -0.001 | -0.001 | -0.002 | -0.002 | -0.002 | -0.002 | 0.0005 | 0.0003 | 0.0006 | 0.0004 |
|  | (0.002) | (0.002) | (0.002) | (0.002) | (0.002) | (0.002) | (0.002) | (0.002) | (0.002) | (0.0017) | (0.002) | (0.002) |
| Intellectual Abilities | $0.441^{* * *}$ | 0.445*** | 0.416*** | $0.421^{* * *}$ | 0.904*** | 0.906*** | 0.900*** | 0.903*** | 0.596*** | 0.599*** | 0.595*** | 0.599*** |
| (log reading score) | (0.122) | (0.120) | (0.121) | (0.120) | (0.134) | (0.134) | (0.134) | (0.134) | (0.099) | (0.099) | (0.098) | (0.099) |
| Fatherố Education | -0.006 | -0.006 | -0.005 | -0.005 | -0.048* | -0.048* | -0.048* | -0.048* | -0.006 | -0.007 | -0.007 | -0.007 |
|  | (0.021) | (0.021) | (0.021) | (0.021) | (0.028) | (0.028) | (0.028) | (0.028) | (0.024) | (0.024) | (0.024) | (0.024) |
| Motherố Education | 0.010 | 0.010 | 0.008 | 0.007 | -0.026 | -0.026 | -0.026 | -0.026 | -0.027 | -0.028 | -0.027 | -0.028 |
|  | (0.022) | (0.022) | (0.021) | (0.022) | (0.028) | (0.028) | (0.028) | (0.028) | (0.021) | (0.021) | (0.021) | (0.021) |
| Economic, Social and Cultural Status | 0.062 | 0.063 | 0.063 | 0.064 | 0.293*** | 0.294*** | 0.294*** | 0.294*** | 0.130*** | $0.131^{* * *}$ | $0.132 * * *$ | 0.132*** |
|  | (0.042) | (0.042) | (0.041) | (0.041) | (0.051) | (0.052) | (0.051) | (0.051) | (0.042) | (0.042) | (0.041) | (0.041) |
| Family Spending on Education | 0.022 | 0.022 | 0.017 | 0.017 | -0.004 | -0.004 | -0.004 | -0.004 | 0.005 | 0.004 | 0.002 | 0.002 |
|  | (0.014) | (0.014) | (0.014) | (0.014) | (0.016) | (0.016) | (0.016) | (0.016) | (0.014) | (0.014) | (0.014) | (0.014) |
| Parental Support for Learning at Home | 0.131*** | 0.131*** | 0.127*** | $0.127^{* * *}$ | 0.101*** | 0.102*** | 0.101*** | 0.102*** | 0.094*** | 0.094*** | 0.093*** | 0.093*** |
|  | (0.018) | (0.018) | (0.018) | (0.018) | (0.024) | (0.024) | (0.024) | (0.024) | (0.018) | (0.018) | (0.018) | (0.018) |
| Parental Emotional Support | -0.011 | -0.012 | -0.009 | -0.009 | 0.019 | 0.019 | 0.019 | 0.019 | -0.014 | -0.014 | -0.013 | -0.013 |
|  | (0.019) | (0.019) | (0.019) | (0.019) | (0.020) | (0.020) | (0.020) | (0.020) | (0.017) | (0.018) | (0.017) | (0.018) |
| Skipping (some) Classes | 0.016 | 0.018 | 0.032 | 0.033 | -0.005 | -0.004 | -0.003 | -0.003 | 0.117 | 0.119 | 0.123 | 0.124 |
|  | (0.082) | (0.08) | (0.083) | (0.084) | (0.133) | (0.133) | (0.132) | (0.133) | (0.087) | (0.088) | (0.085) | (0.087) |
| Coming to School Late | $-0.069^{* *}$ | -0.069** | -0.074** | -0.073** | -0.112** | -0.111** | -0.112** | $-0.112^{* *}$ | -0.099** | $-0.098^{* *}$ | $-0.096 * *$ | -0.096** |
|  | (0.034) | (0.034) | (0.033) | (0.033) | (0.0489 | (0.048) | (0.048) | (0.048) | (0.040) | (0.041) | (0.040) | (0.040) |
| Chatting Online (Outside of School) | 0.008 | 0.010 | 0.010 | 0.011 | 0.005 | 0.006 | 0.005 | 0.006 | 0.022** | 0.024** | 0.023** | 0.024** |
|  | (0.012) | (0.012) | (0.011) | (0.011) | (0.014) | (0.014) | (0.014) | (0.014) | (0.011) | (0.011) | (0.010) | (0.011) |
| Participation in Social Networks | $-0.058 * * *$ | -0.059*** | $-0.056 * * *$ | -0.056*** | -0.039*** | -0.040 *** | -0.039*** | -0.040*** | -0.060 *** | -0.061 *** | -0.059*** | -0.060 *** |
|  | (0.013) | (0.013) | (0.013) | (0.013) | (0.015) | (0.015) | (0.015) | (0.015) | (0.012) | (0.012) | (0.012) | (0.011) |
| Chatting Online (School) | 0.047*** | 0.045*** | 0.045*** | 0.043*** | 0.087*** | 0.086*** | 0.087*** | 0.086*** | 0.061*** | 0.058*** | 0.063*** | 0.060*** |


|  | (0.017) | (0.016) | (0.016) | (0.016) | (0.022) | (0.023) | (0.022) | (0.023) | (0.020) | (0.020) | (0.020) | (0.020) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Feeling Belonging to School | 0.116*** | 0.116*** | 0.105*** | 0.105*** | 0.178*** | 0.177*** | 0.176*** | 0.177*** | 0.127*** | 0.126*** | 0.123*** | 0.123*** |
|  | (0.023) | (0.023) | (0.022) | (0.022) | (0.025) | (0.025) | (0.025) | (0.026) | (0.023) | (0.023) | (0.023) | (0.023) |
| Unfairness of Teacher | -0.003 | -0.003 | -0.004 | -0.004 | 0.008 | 0.008 | 0.008 | 0.008 | -0.010* | -0.010* | -0.011* | -0.011* |
|  | (0.005) | (0.005) | (0.005) | (0.005) | (0.007) | (0.007) | (0.007) | (0.007) | (0.006) | (0.006) | (0.006) | (0.006) |
| Number of Observations | 3,249 | 3,249 | 3,249 | 3,249 | 3,252 | 3,252 | 3,252 | 3,252 | 3,239 | 3,239 | 3,239 | 3,239 |
| Number of Schools | 105 | 105 | 105 | 105 | 105 | 105 | 105 | 105 | 105 | 105 | 105 | 105 |
| $\mathrm{R}^{2}$ | 0.092 | 0.094 |  |  | 0.114 | 0.114 |  |  | 0.118 | 0.120 |  |  |
| Wald Chi2 |  |  | 347.3*** | 399.6*** |  |  | 534.6*** | 547*** |  |  | 623*** | 614.1*** |

Note: Robust standard errors are in parentheses. Robust standard errors are clustered at the school level in the linear estimations. *p<.10, ** $p<.05, * * * p<.001$.

Table 3. Gender and Gender-Matching Effects on Cognitive Performance, public schools

| Dependent Variable | Log Science Score |  |  |  | Log Math Score |  |  |  | Log Reading Score |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Method | OLS |  | Multilevel |  | OLS |  | Multilevel |  | OLS |  | Multilevel |  |
|  | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) | (9) | (10) | (11) | (12) |
| Female Student | -0.001 | -0.002 | -0.007 | 0.001 | -0.004 | -0.021 | -0.011 | -0.025** | 0.051*** | $0.043 * * *$ | $0.043^{* * *}$ | $0.033 * * *$ |
|  | (0.012) | (0.015) | (0.010) | (0.013) | (0.010) | (0.014) | (0.008) | (0.012) | (0.010) | (0.015) | (0.008) | (0.011) |
| Boy School | 0.004 | 0.003 | 0.0003 | 0.002 | 0.007 | 0.004 | 0.002 | -0.001 | 0.007 | 0.006 | -0.001 | -0.003 |
|  | (0.028) | (0.028) | (0.029) | (0.029) | (0.024) | (0.024) | (0.026) | (0.025) | (0.029) | (0.029) | (0.031) | (0.031) |
| Girl School | -0.007 | -0.007 | -0.004 | -0.004 | -0.029 | -0.028 | -0.024 | -0.024 | 0.007 | 0.007 | 0.010 | 0.011 |
|  | (0.019) | (0.019) | (0.022) | (0.022) | (0.024) | (0.023) | (0.028) | (0.027) | (0.017) | (0.017) | (0.020) | (0.020) |
| Community Size | -0.003 | -0.003 | -0.001 | -0.002 | 0.0004 | 0.001 | 0.003 | 0.003 | -0.0005 | -0.0002 | 0.002 | 0.003 |
|  | (0.011) | (0.011) | (0.012) | (0.012) | (0.011) | (0.011) | (0.012) | (0.012) | (0.011) | (0.011) | (0.012) | (0.012) |
| Student-Teacher Ratio | -1.266* | -1.268* | $-1.596 * *$ | -1.574** | -1.163 | -1.221* | $-1.582 * *$ | -1.626** | -0.865 | -0.892 | -1.311 | -1.343 |
|  | (0.725) | (0.724) | (0.793) | (0.799) | (0.737) | (0.736) | (0.796) | (0.794) | (0.763) | (0.765) | (0.826) | (0.828) |
| School Size | 0.00002 | 0.00002 | 0.00003 | 0.00003 | 0.00004 | 0.00004 | 0.00004 | 0.00004 | 0.00003 | 0.00003 | 0.00003 | 0.00003 |
|  | (0.00003) | (0.00003) | (0.00003) | (0.00003) | (0.00003) | (0.00003) | (0.00003) | (0.00003) | (0.00003) | (0.00003) | (0.00004) | (0.00004) |
| Perceived School Quality | 0.006 | 0.006 | -0.001 | -0.001 | 0.007 | 0.007 | -0.0001 | -0.0003 | 0.007 | 0.006 | -0.0009 | -0.001 |
|  | (0.006) | (0.006) | (0.005) | (0.005) | (0.006) | (0.006) | (0.005) | (0.005) | (0.006) | (0.006) | (0.005) | (0.005) |
| Female Teacher | 0.009 | 0.009 | 0.018** | 0.024** | 0.006 | -0.006 | 0.009 | -0.0007 | 0.001 | -0.004 | 0.0009 | -0.006 |
|  | (0.010) | (0.015) | (0.008) | (0.011) | (0.011) | (0.015) | (0.008) | (0.011) | (0.011) | (0.015) | (0.008) | (0.011) |
| Female Student*Female Teacher |  | 0.001 |  | -0.012 |  | 0.027 |  | 0.021 |  | 0.012 |  | 0.015 |
|  |  | (0.015) |  | (0.013) |  | (0.015) |  | (0.015) |  | (0.017) |  | (0.014) |
| Teacherôs Tenure | -0.005 | -0.005 | -0.006 | -0.006 | -0.009 | -0.009 | -0.011 | -0.011 | -0.021 | -0.021 | -0.015 | -0.015 |
|  | (0.013) | (0.013) | (0.010) | (0.010) | (0.015) | (0.015) | (0.011) | (0.011) | (0.014) | (0.014) | (0.011) | (0.011) |
| Experience of Teacher | -0.0001 | -0.0001 | 0.0004 | 0.0004 | -0.0004 | -0.0004 | 9.04e-06 | -0.00002 | -0.0005 | -0.0005 | -0.0004 | -0.0004 |


|  | (0.0005) | (0.0005) | (0.0004) | (0.0004) | (0.0006) | (0.0006) | (0.0005) | (0.0005) | (0.0006) | (0.0006) | (0.0005) | (0.0005) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Fatherố Education | 0.007 | 0.007 | 0.003 | 0.003 | 0.009 | 0.009 | 0.005 | 0.005 | 0.006 | 0.006 | 0.001 | 0.001 |
|  | (0.006) | (0.006) | (0.005) | (0.005) | (0.006) | (0.006) | (0.006) | (0.006) | (0.006) | (0.006) | (0.005) | (0.006) |
| Motherô Education | 0.008 | 0.008 | 0.006 | 0.006 | 0.008 | 0.009* | 0.007 | 0.007 | 0.008 | 0.008 | 0.007 | 0.007 |
|  | (0.006) | (0.006) | (0.005) | (0.005) | (0.005) | (0.005) | (0.005) | (0.005) | (0.005) | (0.005) | (0.005) | (0.005) |
| Economic, Social and Cultural Status | 0.046*** | 0.046*** | 0.033*** | 0.033*** | 0.057*** | $0.057^{* * *}$ | 0.047*** | 0.047*** | 0.057*** | 0.057*** | 0.044*** | 0.044*** |
|  | (0.009) | (0.009) | (0.008) | (0.008) | (0.008) | (0.008) | (0.008) | (0.008) | (0.008) | (0.009) | (0.008) | (0.008) |
| Family Spending on Education | 0.016*** | 0.016*** | 0.009*** | 0.009*** | 0.020*** | $0.020^{* * *}$ | 0.013*** | 0.013*** | 0.019*** | 0.019*** | 0.012*** | 0.012*** |
|  | (0.003) | (0.003) | (0.003) | (0.003) | (0.004) | (0.004) | (0.003) | (0.003) | (0.004) | (0.004) | (0.003) | (0.003) |
| Parental Support for Learning at Home | 0.006 | 0.006 | 0.005 | 0.005 | 0.0002 | 0.00001 | -0.0007 | -0.0009 | 0.005 | 0.005 | 0.004 | 0.003 |
|  | (0.005) | (0.005) | (0.005) | (0.005) | (0.004) | (0.004) | (0.004) | (0.004) | (0.005) | (0.005) | (0.004) | (0.004) |
| Parental Emotional Support | 0.010** | 0.010** | 0.009** | 0.009** | 0.010*** | 0.010*** | 0.009*** | 0.009*** | 0.0077* | 0.008* | 0.007* | 0.007* |
|  | (0.004) | (0.004) | (0.004) | (0.004) | (0.004) | (0.004) | (0.004) | (0.004) | (0.0043) | (0.004) | (0.004) | (0.004) |
| Skipping (some) Classes | $-0.074 * *$ | -0.074** | -0.060** | -0.060** | -0.062** | -0.062** | -0.046** | -0.046** | -0.093*** | -0.093*** | -0.077 *** | -0.077*** |
|  | (0.029) | (0.029) | (0.028) | (0.028) | (0.024) | (0.024) | (0.022) | (0.022) | (0.030) | (0.030) | (0.030) | (0.030) |
| Coming to School Late | $-0.019 * * *$ | -0.019*** | $-0.018^{* * *}$ | -0.018*** | -0.025*** | -0.024*** | $-0.023 * * *$ | -0.023*** | -0.019*** | -0.019*** | $-0.017^{* *}$ | -0.017*** |
|  | (0.006) | (0.006) | (0.006) | (0.006) | (0.006) | (0.006) | (0.006) | (0.006) | (0.007) | (0.007) | (0.007) | (0.007) |
| Chatting Online (Outside of School) | -0.0002 | -0.0001 | 0.001 | 0.001 | 0.001 | 0.001 | 0.002 | 0.002 | -0.0004 | -0.0004 | 0.0007 | 0.0007 |
|  | (0.002) | (0.002) | (0.002) | (0.002) | (0.002) | (0.002) | (0.002) | (0.002) | (0.002) | (0.002) | (0.002) | (0.002) |
| Participation in Social Networks | -0.003 | -0.003 | -0.003 | -0.003 | -0.0004 | -0.0003 | -0.0006 | -0.0006 | -0.0009 | -0.0009 | -0.0008 | -0.0008 |
|  | (0.003) | (0.003) | (0.003) | (0.003) | (0.003) | (0.003) | (0.003) | (0.003) | (0.003) | (0.003) | (0.003) | (0.003) |
| Chatting Online (School) | $-0.028^{* * *}$ | -0.028*** | $-0.024^{* * *}$ | -0.024*** | $-0.027^{* * *}$ | $-0.027 * * *$ | $-0.022^{* * *}$ | -0.021*** | -0.033*** | $-0.033^{* * *}$ | $-0.029 * * *$ | -0.029*** |
|  | (0.005) | (0.005) | (0.004) | (0.004) | (0.005) | (0.005) | (0.004) | (0.004) | (0.005) | (0.005) | (0.005) | (0.005) |
| Feeling Belonging to School | 0.003 | 0.003 | 0.001 | 0.0001 | 0.013*** | 0.013*** | 0.011*** | 0.011*** | 0.002 | 0.002 | 0.0007 | 0.0006 |
|  | (0.005) | (0.005) | (0.004) | (0.004) | (0.005) | (0.005) | (0.004) | (0.004) | (0.005) | (0.005) | (0.004) | (0.004) |
| Unfairness of Teacher | $-0.008 * * *$ | $-0.008 * * *$ | $-0.007 * * *$ | $-0.007 * * *$ | $-0.007^{* * *}$ | $0.007^{* * *}$ | $-0.007 * * *$ | -0.007*** | $-0.008 * * *$ | $-0.008 * * *$ | $-0.007^{* * *}$ | $-0.007 * * *$ |


|  | (0.001) | (0.001) | (0.001) | (0.001) | (0.001) | (0.001) | (0.001) | (0.001) | (0.001) | (0.001) | (0.001) | (0.001) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Number of Observations | 2,280 | 2,280 | 2,280 | 2,280 | 2,280 | 2,280 | 2,280 | 2,280 | 2,280 | 2,280 | 2,280 | 2,280 |
| Number of Schools | 74 | 74 | 74 | 74 | 74 | 74 | 74 | 74 | 74 | 74 | 74 | 74 |
| $\mathrm{R}^{2}$ | 0.217 | 0.217 |  |  | 0.250 | 0.251 |  |  | 0.255 | 0.255 |  |  |
| Wald Chi2 |  |  | 364.1*** | 367.1*** |  |  | 553*** | 568.7*** |  |  | $363.2 * * *$ | 376.1*** |

Note: Robust standard errors are in parentheses. Robust standard errors are clustered at the school level in the linear estimations. $* p<.10, * * p<.05, * * * p<.001$.

Table 4. Gender and Gender-Matching Effects on Non-cognitive Performance, public schools

| Dependent Variable |  | umental M | ation in Scie |  |  | Confiden | n Science |  |  | Interest | Science |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Method | OLS |  | Multilevel |  | OLS |  | Multilevel |  | OLS |  | Multilevel |  |
|  | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) | (9) | (10) | (11) | (12) |
| Female Student | $-0.222 * * *$ | -0.354*** | -0.212*** | $-0.335 * * *$ | 0.023 | -0.031 | 0.022 | -0.030 | $-0.330 * * *$ | $-0.448 * * *$ | -0.326*** | -0.434*** |
|  | (0.056) | (0.145) | (0.054) | (0.072) | (0.051) | (0.098) | (0.050) | (0.097) | (0.043) | (0.068) | (0.040) | (0.066) |
| Boy School | 0.021 | -0.005 | 0.035 | 0.008 | -0.045 | -0.056 | -0.044 | -0.055 | 0.057 | 0.033 | 0.070 | 0.046 |
|  | (0.060) | (0.060) | (0.061) | (0.061) | (0.078) | (0.080) | (0.078) | (0.080) | (0.079) | (0.079) | (0.078) | (0.079) |
| Girl School | 0.193** | 0.193** | 0.187** | 0.188** | -0.154 | -0.154 | -0.153 | -0.153 | 0.068 | 0.068 | 0.067 | 0.067 |
|  | (0.082) | (0.079) | (0.082) | (0.079) | (0.095) | (0.093) | (0.094) | (0.093) | (0.063) | (0.061) | (0.062) | (0.060) |
| Community Size | 0.020 | 0.025 | 0.019 | 0.022 | -0.011 | -0.009 | -0.011 | -0.009 | 0.006 | 0.010 | 0.006 | 0.009 |
|  | (0.030) | (0.029) | (0.030) | (0.029) | (0.037) | (0.037) | (0.037) | (0.037) | (0.032) | (0.031) | (0.032) | (0.031) |
| Student-Teacher Ratio | -2.141 | -2.500 | -2.167 | -2.500 | -2.591 | -2.738 | -2.624 | -2.759 | -1.265 | -1.599 | -1.260 | -1.567 |
|  | (2.117) | (2.072) | (2.101) | (2.047) | (2.045) | (2.044) | (2.045) | (2.040) | (1.931) | (1.913) | (1.974) | (1.937) |
| School Size | -0.0001 | -0.0001 | -0.00008 | -0.0001 | -5.79e-06 | -8.94e-06 | -5.52e-06 | -8.66e-06 | -0.00007 | -0.00008 | -0.00007 | -0.00007 |
|  | (0.0001) | (0.0001) | (0.0001) | (0.0001) | (0.0001) | (0.0001) | (0.001) | (0.0001) | (0.0001) | (0.0001) | (0.0001) | (0.0001) |
| Perceived School Quality | 0.070*** | 0.069*** | 0.071*** | 0.070*** | 0.070** | 0.070** | 0.070** | 0.069** | 0.082*** | 0.081*** | 0.083*** | $0.082^{* * *}$ |
|  | (0.025) | (0.025) | (0.025) | (0.025) | (0.032) | (0.032) | (0.032) | (0.032) | (0.023) | (0.023) | (0.023) | (0.023) |
| Female Teacher | -0.053 | -0.145*** | -0.036 | -0.126** | -0.113** | -0.151* | -0.112** | -0.149* | -0.013 | -0.096 | 0.004 | -0.075 |
|  | (0.045) | (0.055) | (0.044) | (0.054) | (0.052) | (0.078) | (0.051) | (0.077) | (0.044) | (0.062) | (0.043) | (0.061) |
| Female Student*Female Teacher |  | 0.202** |  | 0.186** |  | 0.081 |  | 0.079 |  | 0.180** |  | 0.163** |
|  |  | (0.079) |  | (0.077) |  | (0.116) |  | (0.115) |  | (0.084) |  | (0.061) |
| Teacherôs Tenure | -0.016 | -0.015 | -0.046 | -0.043 | 0.024 | 0.024 | 0.025 | 0.025 | -0.065 | -0.064 | -0.063 | -0.063 |
|  | (0.002) | (0.064) | (0.063) | (0.064) | (0.070) | (0.070) | (0.070) | (0.070) | (0.056) | (0.056) | (0.057) | (0.056) |
| Experience of Teacher | -0.002 | -0.002 | -0.001 | -0.001 | -0.001 | -0.001 | -0.001 | -0.001 | 0.0004 | 0.0001 | 0.0004 | 0.0002 |


|  | (0.002) | (0.002) | (0.002) | (0.002) | (0.003) | (0.003) | (0.003) | (0.003) | (0.002) | (0.002) | (0.002) | (0.002) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Intellectual Abilities | 0.375*** | 0.379*** | $0.359 * * *$ | 0.364*** | 0.819*** | 0.820 *** | 0.815*** | 0.817*** | 0.629*** | $0.631^{* * *}$ | $0.623^{* * *}$ | 0.626*** |
| (log reading score) | (0.137) | (0.135) | (0.137) | (0.135) | (0.136) | (0.135) | (0.135) | (0.135) | (0.120) | (0.120) | (0.118) | (0.118) |
| Fatherố Education | -0.022 | -0.022 | -0.022 | -0.022 | -0.032 | -0.032 | -0.032 | -0.032 | -0.001 | -0.001 | -0.002 | -0.001 |
|  | (0.020) | (0.020) | (0.020) | (0.020) | (0.029) | (0.029) | (0.030) | (0.029) | (0.029) | (0.028) | (0.028) | (0.028) |
| Motherôs Education | 0.003 | -0.0001 | -0.0001 | -0.002 | $-0.069^{* *}$ | -0.070** | $-0.069 * *$ | -0.070** | -0.039* | -0.041* | -0.039* | -0.042* |
|  | (0.025) | (0.024) | (0.025) | (0.024) | (0.034) | (0.034) | (0.033) | (0.034) | (0.022) | (0.022) | (0.022) | (0.022) |
| Economic, Social and Cultural Status | 0.067 | 0.067 | 0.072 | 0.071 | 0.328*** | $0.328 * * *$ | 0.328*** | 0.328*** | $0.141^{* * *}$ | 0.141*** | 0.146*** | 0.146*** |
|  | (0.047) | (0.047) | (0.047) | (0.047) | (0.060) | (0.060) | (0.059) | (0.059) | (0.048) | (0.022) | (0.047) | (0.047) |
| Family Spending on Education | 0.027* | 0.026* | 0.022 | 0.022 | -0.005 | -0.005 | -0.005 | -0.005 | 0.020 | 0.020 | 0.019 | 0.019 |
|  | (0.014) | (0.015) | (0.014) | (0.014) | (0.018) | (0.018) | (0.018) | (0.018) | (0.018) | (0.018) | (0.017) | (0.017) |
| Parental Support for Learning at Home | 0.118*** | $0.119 * * *$ | 0.115*** | $0.116^{* * *}$ | 0.071** | 0.071** | 0.071** | 0.071** | 0.075*** | 0.077*** | 0.074*** | 0.075*** |
|  | (0.021) | (0.021) | (0.021) | (0.021) | (0.028) | (0.029) | (0.028) | (0.028) | (0.022) | (0.022) | (0.022) | (0.022) |
| Parental Emotional Support | 0.011 | 0.011 | 0.011 | 0.011 | 0.022 | 0.022 | 0.022 | 0.022 | -0.017 | -0.017 | -0.017 | -0.017 |
|  | (0.023) | (0.023) | (0.023) | (0.023) | (0.025) | (0.025) | (0.024) | (0.025) | (0.020) | (0.021) | (0.020) | (0.020) |
| Skipping (some) Classes | 0.026 | 0.026 | 0.032 | 0.032 | 0.009 | 0.010 | 0.011 | 0.011 | 0.132 | 0.132 | 0.132 | 0.133 |
|  | (0.091) | (0.093) | (0.093) | (0.093) | (0.146) | (0.146) | (0.015) | (0.146) | (0.095) | (0.096) | (0.093) | (0.095) |
| Coming to School Late | -0.069* | -0.067* | -0.067* | -0.065* | $-0.121^{* *}$ | -0.120** | $-0.121^{* *}$ | -0.120** | $-0.119 * * *$ | $-0.116^{* * *}$ | $-0.112 * * *$ | $-0.110^{* * *}$ |
|  | (0.038) | (0.039) | (0.037) | (0.038) | (0.055) | (0.055) | (0.054) | (0.054) | (0.043) | (0.043) | (0.042) | (0.043) |
| Chatting Online (Outside of School) | 0.018 | 0.019 | 0.018 | 0.019 | -0.008 | -0.007 | -0.008 | -0.007 | 0.031** | 0.032*** | 0.031*** | 0.032*** |
|  | (0.013) | (0.012) | (0.012) | (0.012) | (0.017) | (0.017) | (0.017) | (0.017) | (0.012) | (0.012) | (0.012) | (0.012) |
| Participation in Social Networks | $-0.053 * * *$ | $-0.055 * * *$ | $-0.054 * * *$ | -0.055*** | -0.050*** | -0.051** | -0.050 *** | $-0.051 * * *$ | $-0.058 * * *$ | $-0.059 * * *$ | $-0.058^{* * *}$ | $-0.059 * * *$ |
|  | (0.015) | (0.015) | (0.014) | (0.014) | (0.018) | (0.018) | (0.017) | (0.018) | (0.013) | (0.013) | (0.013) | (0.013) |
| Chatting Online (School) | 0.030* | 0.029 | 0.033* | 0.032* | 0.104*** | 0.104*** | 0.104*** | 0.104*** | 0.071*** | 0.070*** | 0.072*** | 0.071*** |
|  | (0.018) | (0.018) | (0.017) | (0.017) | (0.027) | (0.027) | (0.026) | (0.027) | (0.023) | (0.023) | (0.023) | (0.023) |
| Feeling Belonging to School | $0.119^{* * *}$ | 0.119*** | 0.113*** | $0.113^{* * *}$ | 0.190*** | $0.190^{* * *}$ | 0.189*** | 0.189*** | 0.131*** | 0.131*** | $0.128^{* * *}$ | 0.129*** |


|  | (0.026) | (0.026) | (0.026) | (0.026) | (0.032) | (0.033) | (0.032) | (0.032) | (0.028) | (0.028) | (0.028) | (0.028) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Unfairness of Teacher | -0.004 | -0.004 | -0.004 | -0.004 | 0.010 | 0.010 | 0.010 | 0.010 | -0.016** | -0.016** | -0.017** | -0.016** |
|  | (0.006) | (0.006) | (0.006) | (0.006) | (0.008) | (0.008) | (0.008) | (0.008) | (0.007) | (0.007) | (0.007) | (0.007) |
| Number of Observations | 2,273 | 2,273 | 2,273 | 2,273 | 2,276 | 2,276 | 2,276 | 2,276 | 2,265 | 2,265 | 2,265 | 2,265 |
| Number of Schools | 74 | 74 | 74 | 74 | 74 | 74 | 74 | 74 | 74 | 74 | 74 | 74 |
| $\mathrm{R}^{2}$ | 0.083 | 0.085 |  |  | 0.106 | 0.106 |  |  | 0.117 | 0.119 |  |  |
| Wald Chi2 |  |  | 328.7*** | 378.5*** |  |  | 390.6*** | 403.2*** |  |  | 456.4*** | 467.9*** |

Note: Robust standard errors are in parentheses. Robust standard errors are clustered at the school level in the linear estimations. $* p<.10, * * p<.05, * * * p<.001$.

Table 5. Average Treatment Effects (ATE) of Single-sex Schooling on Cognitive and Non-cognitive Performance, propensity-score matching

Table 5.1. (including both public and private schools)

| Dependent Variable | Sample | ATE (single-sex school) | AI Robust Std.Err. | Number of Obs. |
| :---: | :---: | :---: | :---: | :---: |
| $(\log )$ Science Score | Boys | $0.035^{* * *}$ | 0.013 | 1,793 |
| $(\log )$ Science Score | Girls | -0.010 | 0.013 | 1,485 |
| $(l o g)$ Math Score | Boys | 0.031 | 0.013 | 1,786 |
| $(l o g)$ Math Score | Girls | $-0.021^{*}$ | 0.012 | 1,492 |
| $(\log )$ Reading Score | Boys | $0.039^{* * *}$ | 0.014 | 1,786 |
| (log) Reading Score | Girls | 0.010 | 0.010 | 1,492 |
| Instrumental Motivation | Boys | 0.039 | 0.062 | 1,786 |
| Instrumental Motivation | Girls | -0.008 | 0.069 | 1,482 |
| Confidence in Science | Boys | 0.033 | 0.069 | 1,787 |
| Confidence in Science | Girls | -0.105 | 0.076 | 1,484 |
| Interest in Science | Boys | 0.021 | 0.068 | 1,776 |
| Interest in Science | Girls | 0.021 | 0.069 | 1,482 |

Table 5.2. (public schools only)

| Dependent Variable | Sample | ATE (single-sex school) | AI Robust Std.Err. | Number of Obs. |
| :---: | :---: | :---: | :---: | :---: |
| $(\log )$ Science Score | Boys | 0.019 | 0.018 | 1,256 |
| $(\log )$ Science Score | Girls | -0.094 | 0.071 | 1,039 |
| $(\log )$ Math Score | Boys | 0.016 | 0.015 | 1,237 |
| $(\log )$ Math Score | Girls | -0.011 | 0.052 | 1,025 |
| (log) Reading Score | Boys | 0.013 | 0.017 | 1,237 |
| (log) Reading Score | Girls | -0.012 | 0.070 | 1,025 |
| Instrumental Motivation | Boys | -0.039 | 0.071 | 1,252 |
| Instrumental Motivation | Girls | 0.250 | 0.202 | 1,036 |
| Confidence in Science | Boys | -0.046 | 0.089 | 1,253 |
| Confidence in Science | Girls | $0.430^{*}$ | 0.244 | 1,038 |
| Interest in Science | Boys | 0.103 | 0.079 | 1,244 |
| Interest in Science | Girls | 0.007 | 0.075 | 1,036 |

Table 6. Gender and Gender-Matching Effects on Non-cognitive Performance, heterogeneous responses by quartiles
OLS, including both public and private schools

| Dependent Variable | Instrumental Motivation in Science |  |  |  | Confidence in Science |  |  |  | Interest in Science |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Quartile <br> (Science Scores) | $\begin{gathered} 4^{\text {th }} \\ (>582) \end{gathered}$ | $\begin{gathered} 3^{\mathrm{rd}} \\ (518 \mathrm{i} 582) \end{gathered}$ | $\begin{gathered} 2^{\mathrm{nd}} \\ (449 \mathrm{i} 518) \end{gathered}$ | $1^{s t}$ <br> (<449) | $\begin{gathered} 4^{\text {th }} \\ (>582) \end{gathered}$ | $\begin{gathered} 3^{\text {rd }} \\ (518 \text { ī } 582) \end{gathered}$ | $\begin{gathered} 2^{\text {nd }} \\ (449 \overline{1} 518) \end{gathered}$ | $1^{\text {st }}$ <br> (<449) | $\begin{gathered} 4^{\text {th }} \\ (>582) \end{gathered}$ | $\begin{gathered} 3^{\text {rd }} \\ (518 \mathrm{i} 582) \end{gathered}$ | $\begin{gathered} 2^{\text {nd }} \\ (449 \mathrm{i} 518) \end{gathered}$ | $\begin{gathered} 1^{\text {st }} \\ (<449) \end{gathered}$ |
|  | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) | (9) | (10) | (11) | (12) |
| Female Student | $\begin{gathered} -0.494 * * * \\ (0.102) \end{gathered}$ | $\begin{aligned} & -0.209^{*} \\ & (0.122) \end{aligned}$ | $-0.220^{*}$ <br> (0.117) | $\begin{gathered} -0.234^{*} * \\ (0.112) \end{gathered}$ | $\begin{aligned} & -0.145 \\ & (0.119) \end{aligned}$ | $\begin{gathered} 0.070 \\ (0.113) \end{gathered}$ | $\begin{gathered} 0.111 \\ (0.176) \end{gathered}$ | $\begin{gathered} 0.171 \\ (0.182) \end{gathered}$ | $\begin{gathered} -0.474 * * * \\ (0.095) \end{gathered}$ | $\begin{gathered} -0.229^{* *} \\ (0.109) \end{gathered}$ | $\begin{gathered} -0.495^{* * *} \\ (0.113) \end{gathered}$ | $\begin{gathered} -0.395^{* * *} \\ (0.113) \end{gathered}$ |
| Boy School | -0.045 | 0.004 | -0.013 | 0.051 | -0.257 ** | -0.214* | 0.204 | 0.145 | -0.082 | 0.005 | 0.047 | 0.034 |
|  | (0.093) | (0.100) | (0.134) | (0.103) | (0.100) | (0.112) | (0.180) | (0.148) | (0.073) | (0.086) | (0.109) | (0.111) |
| Girl School | 0.202 | 0.014 | 0.154 | 0.063 | 0.110 | -0.160 | -0.413** | 0.121 | 0.176 | -0.034 | 0.088 | 0.089 |
|  | (0.204) | (0.095) | (0.114) |  | (0.093) | (0.112) | (0.161) | (0.148) | (0.131) | (0.085) | (0.088) | (0.108) |
| Female Teacher | -0.243** | -0.177* | -0.040 | -0.107 | -0.253** | 0.006 | 0.0001 | -0.133 | -0.060 | -0.043 | -0.184 | -0.022 |
|  | (0.099) | (0.090) | (0.111) | (0.096) | (0.100) | (0.100) | (0.148) | (0.145) | (0.075) | (0.107) | (0.126) | (0.105) |
| Female Student*Female Teacher | 0.373** | 0.091 | -0.018 | 0.254* | 0.199 | -0.058 | -0.001 | 0.076 | 0.222** | -0.071 | 0.390*** | 0.154 |
|  | (0.145) | (0.125) | (0.111) | (0.139) | (0.150) | (0.133) | (0.199) | (0.218) | (0.107) | (0.134) | (0.141) | (0.150) |
| School Inputs | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| Teacherôs Inputs | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| Family Inputs | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| Behavioral Factors | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| Relational Factors | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| Number of Observations | 855 | 832 | 775 | 787 | 855 | 833 | 774 | 790 | 853 | 832 | 772 | 782 |
| Number of Schools | 98 | 104 | 104 | 97 | 98 | 104 | 104 | 98 | 98 | 104 | 104 | 98 |
| $\mathrm{R}^{2}$ | 0.122 | 0.090 | 0.077 | 0.070 | 0.120 | 0.080 | 0.125 | 0.072 | 0.152 | 0.115 | 0.087 | 0.088 |

Note: Robust standard errors are in parentheses. Parentheses are robust standard errors clustered at the school level. $* p<.10, * * p<.05, * * * p<.001$.

Table 7. Gender and Gender-Matching Effects on Non-cognitive Performance, heterogeneous responses by quartiles
OLS, public schools only


Note: Robust standard errors are in parentheses. Parentheses are robust standard errors clustered at the school level. $* p<.10, * * p<.05, * * * p<.001$.

Table 8. Average Treatment Effects (ATE) of Single-sex Schooling on Non-cognitive
Performance, heterogeneous responses by quartiles, propensity-score matching

Table 8.1. (including both public and private schools)

| DV | Instrumental Motivation in Science |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Quartile | $4^{\text {th }}$ ( score > 582) |  | $3^{\text {rd }}$ ( 518 < score < 582) |  | $2^{\text {nd }}$ (449 < score < 518) |  | $1^{\text {st }}$ (score < 449) |  |
| Gender of Students | Boys | Girls | Boys | Girls | Boys | Girls | Boys | Girls |
| ATE | -0.044 | 0.443*** | -0.048 | -0.129 | 0.056 | 0.050 | 0.062 | 0.220** |
| AI Robust Std.Err. | 0.069 | 0.116 | 0.115 | 0.124 | 0.093 | 0.146 | 0.141 | 0.101 |
| Number of Obs. | 496 | 365 | 427 | 407 | 390 | 394 | 474 | 320 |
| DV | Confidence in Science |  |  |  |  |  |  |  |
| Quartile | $4^{\text {th }}$ ( score $>582$ ) |  | $3^{\text {rd }}$ ( 518 < score < 582) |  | $2^{\text {nd }}$ (449 < score < 518) |  | $1^{\text {st }}$ (score < 449) |  |
| Gender of Students | Boys | Girls | Boys | Girls | Boys | Girls | Boys | Girls |
| ATE | -0.319*** | 0.154* | -0.233** | -0.690*** | 0.302 | -0.367*** | 0.072 | 0.235* |
| AI Robust Std.Err. | 0.096 | 0.079 | 0.095 | 0.224 | 0.149 | 0.135 | 0.152 | 0.120 |
| Number of Obs. |  | 365 |  |  | 389 | 394 | 476 | 321 |
| DV | Interest in Science |  |  |  |  |  |  |  |
| Quartile | $4^{\text {th }}$ (score > 582) |  | $3^{\text {rd }}$ (518 < score < 582) |  | $2^{\text {nd }}$ (449 < score < 518) |  | $1^{\text {st }}($ score < 449) |  |
| Gender of Students | Boys | Girls | Boys | Girls | Boys | Girls | Boys | Girls |
| ATE | -0.069 | 0.235*** | -0.068 | -0.325** | 0.291 | 0.249 | 0.139 | 0.053 |
| AI Robust Std.Err. | 0.044 | 0.062 | 0.103 | 0.138 | 0.194 | 0.189 | 0.120 | 0.110 |
| Number of Obs. | 494 | 365 | 426 | 408 | 389 | 392 | 468 | 321 |

Table 8.2. (public schools only)

| DV | Instrumental Motivation in Science |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Quartile | $4^{\text {th }}$ ( score > 582) |  | $3^{\text {rd }}$ (518 < score < 582) |  | $2^{\text {nd }}$ (449 < score < 518) |  | $1^{\text {st }}$ (score < 449) |  |
| Gender of Students | Boys | Girls | Boys | Girls | Boys | Girls | Boys | Girls |
| ATE | -0.017 | 0.642*** | -0.123 | 0.127 | 0.207 | 0.220 | -0.047 | -0.048 |
| AI Robust Std.Err. | 0.111 | 0.157 | 0.119 | 0.103 | 0.259 | 0.619 | 0.234 | 0.136 |
| Number of Obs. | 292 | 223 | 300 | 282 | 293 | 292 | 368 | 232 |
| DV | Confidence in Science |  |  |  |  |  |  |  |
| Quartile | $4^{\text {th }}$ (score > 582) |  | $3{ }^{\text {rd }}$ ( 518 < score < 582 ) |  | $2^{\text {nd }}$ (449 < score < 518) |  | $1^{\text {st }}$ (score < 449) |  |
| Gender of Students | Boys | Girls | Boys | Girls | Boys | Girls | Boys | Girls |
| ATE | -0.335** | 0.402** | -0.382** | -0.014 | 0.348** | -0.090 | 0.222 | 0.047 |
| AI Robust Std.Err. | 0.142 | 0.198 | 0.158 | 0.106 | 0.167 | 0.209 | 0.235 | 0.121 |
| Number of Obs. | 292 | 223 | 300 | 283 | 292 | 292 | 370 | 243 |
| DV | Interest in Science |  |  |  |  |  |  |  |
| Quartile | $4^{\text {th }}$ (score $>582$ ) |  | $3^{\text {rd }}(518<\text { score }<582)$ |  | $2^{\text {nd }}(449<\text { score }<518)$ |  | $1^{\text {st }}$ (score < 449) |  |
| Gender of Students | Boys | Girls | Boys | Girls | Boys | Girls | Boys | Girls |
| ATE | -0.050 | 0.330** | 0.039 | -0.192** | 0.021 | -0.081 | 0.226 | -0.607 |
| AI Robust Std.Err. | 0.128 | 0.129 | 0.133 | 0.091 | 0.212 | 0.179 | 0.151 | 0.516 |
| Number of Obs. | 291 | 223 | 299 | 283 | 292 | 290 | 363 | 243 |

Figure 1. Gender Differences in Cognitive and Non-cognitive Performance in South Korea (PISA 2015, OECD)
(a) Science Score
(b) Math Score
(c) Reading Score



(d) Instrumental Motivation in Science (index)
(e) Interest in Science (index)


Figure 2. Quantile Distribution of Cognitive and Non-Cognitive Performance, by gender
(1) Science Scores

(2) Math Scores


Boys $(\mathrm{n}=1,793)$


Girls ( $\mathrm{n}=1,466$ )
(3) Reading Scores


Boys ( $\mathrm{n}=1,793$ )


Girls ( $\mathrm{n}=1,466$ )
(4) Instrumental Motivation in Science

(5) Confidence in Science

(6) Interest in Science


Appendix A. Descriptive Statistics

| Variable | Observations | Mean | Std. Dev. | Min. | Max. |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Science Score | 3,259 | 517.95 | 96.77 | 192.38 | 788.37 |
| Math Score | 3,259 | 526.29 | 101.52 | 132.19 | 827.77 |
| Reading Score | 3,259 | 517.05 | 99.24 | 148.47 | 804.33 |
| Instrumental Motivation (index) | 3,249 | 0.03 | 1.01 | -1.93 | 1.74 |
| Confidence in Science (index) | 3,252 | -0.01 | 1.22 | -3.76 | 3.28 |
| Interest in Science (index) | 3,239 | -0.07 | 0.99 | -2.55 | 2.56 |
| Female Student (dummy) | 3,259 | 0.45 | 0.50 | 0 | 1 |
| Public School (dummy) | 3,259 | 0.70 | 0.46 | 0 | 1 |
| Community Size (index) | 3,259 | 4.27 | 0.85 | 1 | 5 |
| Student-Teacher Ratio | 3,259 | 14.32 | 2.57 | 7.2 | 20.83 |
| School Size | 3,259 | 989.81 | 343.68 | 72 | 1,679 |
| Perceived School Quality (index) | 3,259 | -0.05 | 0.867 | -3.55 | 2.53 |
| Female Teacher (science, dummy) | 3,259 | 0.52 | 0.50 | 0 | 1 |
| Female Teacher (main, dummy) | 3,220 | 0.53 | 0.50 | 0 | 1 |
| Teacherố Tenure (science, dummy) | 3,259 | 0.83 | 0.38 | 0 | 1 |
| Teacherô Tenure (main, dummy) | 3,224 | 0.83 | 0.38 | 0 | 1 |
| Teacherô Experience (science) | 3,259 | 16.38 | 10.05 | 0 | 40 |
| Teacherô Experience (main) | 3,209 | 16.42 | 10.04 | 0 | 40 |
| Fatherô Education (index) | 3,259 | 5.38 | 1.01 | 1 | 7 |
| Motherô Education (index) | 3,259 | 5.24 | 0.99 | 1 | 7 |
| Feeling Belonging to School (index) | 3,259 | 0.14 | 0.86 | -3.13 | 2.59 |
| Unfairness of Teacher (index) | 3,259 | 8.34 | 3.14 | 2 | 24 |
| Economic, Social and Cultural Status (index) | 3,259 | -0.19 | 0.69 | -4.08 | 1.91 |
| Family Spending on Education (index) | 3,259 | 3.34 | 1.37 | 1 | 6 |
| Parental Support for Learning at Home (index) | 3,259 | -0.58 | 1.01 | -5.01 | 3.74 |
| Parental Emotional Support (index) | 3,259 | -0.72 | 1.11 | -3.82 | 0.75 |
| Skipping (some) Classes (index) | 3,259 | 1.03 | 0.23 | 1 | 4 |
| Coming to School Late (index) | 3,259 | 1.24 | 0.59 | 1 | 4 |
| Chatting Online (outside of school, index) | 3,259 | 2.64 | 1.66 | 1 | 5 |
| Participation in Social Networks (index) | 3,259 | 3.71 | 1.46 | 1 | 5 |
| Chatting Online (school, index) | 3,259 | 1.45 | 0.98 | 1 | 5 |

## Appendix B. Differences in Performance by Gender

## B.1. Mean Differences

|  | Observations | Mean | Std.Dev. | Min. | Max. |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Science Scores (boys) | 1,793 | 516.46 | 102.43 | 192.38 | 825.03 |
| Science Scores (girl) | 1,466 | 519.78 | 89.36 | 243.14 | 757.16 |
| T-statistics (Bï G) |  |  | -0.98 |  |  |
| Math Scores (boys) | 1,793 | 526.46 | 106.60 | 132.19 | 827.77 |
| Math Scores (girls) | 1,466 | 526.08 | 94.97 | 243.73 | 825.03 |
| T-statistics (Bï G) |  |  | 0.10 |  |  |
| Reading Scores (boys) | 1,793 | 501.98 | 102.73 | 148.47 | 804.33 |
| Reading Scores (girls) | 1,466 | 535.48 | 91.51 | 224.77 | 779.75 |
| T-statistics (Bï G) |  |  | -9.73*** |  |  |
| Instrumental Motivation (boys) | 1,786 | 0.13 | 1.00 | -1.93 | 1.74 |
| Instrumental Motivation (girls) | 1,463 | -0.10 | 1.01 | -1.93 | 1.74 |
| T-statistics (Bï G) |  |  | 6.66*** |  |  |
| Confidence in Science (boys) | 1,787 | -0.03 | 1.29 | -3.76 | 3.28 |
| Confidence in Science (girls) | 1,465 | 0.01 | 1.13 | -3.76 | 3.28 |
| T-statistics (Bï G) |  |  | -1.09 |  |  |
| Interest in Science (boys) | 1,776 | 0.09 | 1.02 | -2.53 | 2.45 |
| Interest in Science (girls) | 1,463 | -0.25 | 0.91 | -2.55 | 2.56 |
| T-statistics (Bï G) |  |  | 9.91*** |  |  |

## B.2. Non-parametric Equality-of-medians test

| Science Scores | Boys | Girls | Total |
| :---: | :---: | :---: | :---: |
| Lower than the Median | 899 | 731 | 1,630 |
| Higher than the Median | 894 | 735 | 1,629 |
| Total | 1,793 | 1,466 | 3,259 |
| Corrected Pearson Chi2(1) |  | 0.01 |  |
| Math Scores | Boys | Girls | Total |
| Lower than the Median | 894 | 736 | 1,630 |
| Higher than the Median | 899 | 730 | 1,629 |
| Total | 1,793 | 1,466 | 3,259 |
| Corrected Pearson Chi2(1) |  | 0.03 |  |
| Reading Scores | Boys | Girls | Total |
| Lower than the Median | 1,002 | 628 | 1,630 |
| Higher than the Median | 791 | 838 | 1,629 |
| Total | 1,793 | 1,466 | 3,259 |
| Corrected Pearson Chi2(1) |  | 54.39*** |  |
| Instrumental Motivation | Boys | Girls | Total |
| Lower than the Median | 1,374 | 1,200 | 2,574 |
| Higher than the Median | 412 | 263 | 675 |
| Total | 1,786 | 1,463 | 3,249 |
| Corrected Pearson Chi2(1) |  | 12.36*** |  |
| Confidence in Science | Boys | Girls | Total |
| Lower than the Median | 899 | 746 | 1,645 |
| Higher than the Median | 888 | 719 | 1,607 |
| Total | 1,787 | 1,465 |  |
| Corrected Pearson Chi2(1) |  | 0.10 |  |
| Interest in Science | Boys | Girls | Total |
| Lower than the Median | 728 | 896 | 1,624 |
| Higher than the Median | 1,048 | 567 | 1,615 |
| Total | 1,776 | 1,463 | 3,239 |
| Corrected Pearson Chi2(1) |  | 130.81*** |  |

## B.3. Two-sample Wilcoxon Rank-sum (Mann-Whitney) Distribution Test

| Science Scores | Observations | Rank Sum | Expected |
| :---: | :---: | :---: | :---: |
| Boys | 1,793 | 2904255 | 2922590 |
| Girls | 1,466 | 2407915 | 2389580 |
| Combined | 3,259 | 5312170 | 5312170 |
| Z-statistics |  | -0.69 |  |
| Math Scores | Observations | Rank Sum | Expected |
| Boys | 1,793 | 2929408.5 | 2922590 |
| Girls | 1,466 | 2382761.5 | 2389580 |
| Combined | 3,259 | 5312170 | 5312170 |
| Z-statistics |  | 0.26 |  |
| Reading Scores | Observations | Rank Sum | Expected |
| Boys | 1,793 | 2680827.5 | 2922590 |
| Girls | 1,466 | 2631342.5 | 2389580 |
| Combined | 3,259 | 5312170 | 5312170 |
| Z-statistics |  | -9.05*** |  |
| Instrumental Motivation in Science | Observations | Rank Sum | Expected |
| Boys | 1,786 | 3094890 | 2902250 |
| Girls | 1,463 | 2184735 | 2377375 |
| Combined | 3,249 | 5279625 | 5279645 |
| Z-statistics |  | 7.36*** |  |
| Confidence in Science | Observations | Rank Sum | Expected |
| Boys | 1,787 | 2865384.5 | 2906555.5 |
| Girls | 1,465 | 2423993.5 | 2382822.5 |
| Combined | 3,252 | 5289378 | 5289378 |
| Z-statistics |  | -1.55 |  |
| Interest in Science | Observations | Rank Sum | Expected |
| Boys | 1,776 | 3193867 | 2877120 |
| Girls | 1,463 | 2053313 | 2370060 |
| Combined | 3,239 | 5247180 | 5247180 |
| Z-statistics |  | 11.97*** |  |

## Appendix C. Determinants of Attending a Single-sex School, probit analysis

|  | Attending an All-Girls School |  | Attending an All-Boys School |  |
| :---: | :---: | :---: | :---: | :---: |
|  | (1) | (2) | (3) | (4) |
| Economic, Social and Cultural Status | -0.172*** | 0.066 | 0.078 | 0.090 |
|  | (0.058) | (0.087) | (0.051) | (0.075) |
| Family Spending on Education | -0.003 | 0.008 | 0.031 | 0.033 |
|  | (0.025) | (0.025) | (0.022) | (0.022) |
| Parental Support for Learning at Home | -0.046 | -0.044 | -0.012 | -0.014 |
|  | (0.035) | (0.035) | (0.029) | (0.030) |
| Parental Emotional Support | -0.038 | -0.032 | 0.092*** | 0.088*** |
|  | (0.030) | (0.030) | (0.028) | (0.028) |
| Community Size | -0.004 | -0.008 | $0.395^{* * *}$ | 0.395*** |
|  | (0.034) | (0.034) | (0.042) | (0.043) |
| Motherô Education | -0.005 | 0.020 | 0.027 | 0.024 |
|  | (0.038) | (0.040) | (0.033) | (0.035) |
| Fatherôs Education |  | $-0.120^{* * *}$ |  | 0.001 |
|  |  | (0.043) |  | (0.037) |
| Cultural Possession at Home |  | -0.016 |  | 0.022 |
|  |  | (0.059) |  | (0.055) |
| Home Educational Resources |  | 0.110** |  | 0.060 |
|  |  | (0.045) |  | (0.039) |
| Home Possessions |  | -0.368** |  | -0.103 |
|  |  | (0.160) |  | (0.138) |
| Family Wealth |  | 0.211 |  | 0.041 |
|  |  | (0.138) |  | (0.115) |
| Number of Observations | 1,980 | 1,965 | 2,334 | 2,302 |
| Log pseudo likelihood | -1,103.05 | -1,082.30 | -1,344.00 | -1,325.80 |

Note: Robust standard errors are in parentheses. Parentheses are robust standard errors. * $p<.10, * * p<.05, * * *$ $p<.001$.


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[^1]:    ${ }^{1}$ Jackson (2012) finds an overall null effect of attending a single-sex school, however, his study further reveals that single-sex schools provide sizeable advantages for girls with strong preferences for attending an all-girls school.

[^2]:    ${ }^{2}$ While single-sex schooling can capture gender-matching effects at the school level, gender-matching between a female teacher and a female student accounts for interaction effects at the individual level.

[^3]:    ${ }^{3}$ In public schools, teachers have to pass the national teacher exam (implemented at the provincial level) to be employed, but this exam is not required for teachers in private schools. Also, teachers in public schools are rotated to different schools regularly (e.g. every five years), while teachers in private schools are not subject to obligatory relocation.

[^4]:    ${ }^{4}$ The mean reading score of girls is 34 points (a third of its standard deviation) higher than that of boys. Furthermore, 57 percent of girls have a reading score above the median, while only 44 percent of boys do. Also, the MannWhitney rank-sum test places more girls on the right-side of the distribution curve (i.e. higher scores) than boys. ${ }^{5}$ On the other hand, there is no difference in the level of confidence between boys and girls.

[^5]:    ${ }^{6}$ Teachersôeducation is not included as a teacherô input because of little variations in the variable. In South Korea, nearly all teachers are certified with a bachelor degree or higher.

[^6]:    ${ }^{7}$ In Tables 6 and 7, the results of the control variables (school, teacher $\hat{\hat{O}}$, family, behavioral, and relational inputs) are not presented to save space but can be obtained from the author upon request.

[^7]:    ${ }^{8}$ ̄ $0.494 \overline{1} 0.243+0.373+0.443=+0.079$ (computed based on Column 1 of Table 6 and Table 8.1).

